

Lecture 1 (21.2.11): Simple random walk (sometimes Markov chain or general random walk)  
← also  $\tau_0, \tau_n, n$  are stop. times.  
Following Lalley: Gambler's ruin & time to hit. Stopping time & Strong Markov Property.  
conclude recurrence of SRW (also by summing return prob.).

First passage time dist. using generating functions.  $P(\tau = 2n-1) = P(S_{2n-1} = 1) / (2n-1)$ .  
Relation to stable law of index  $1/2$ .

Derive dist. of maximum using reflection principle. Use to give another deriv. of first passage dist. ← Catalan number & ballot thm.

Following Lalley: For general skip-free RW, use cyclic perm. trick (start at first max.)  
to show that  $P(\tau = n) = \frac{1}{n} P(S_n = 1)$ . Maybe in exercise, conn. to Lagrange inv.

Wald identities: 3 identities & proofs. Use to get gambler's ruin and time again.  
Possible to get first passage time dist. gen. fcn. using 3rd identity.  
Possible to get  $E S_T | S_T = 0$ , say, for  $T$  hitting time of  $\{-A, B\}$  using third identity.

Lalley also has proof of recurrence of any mean zero 1D RW using Birkhoff's ergodic thm. and range thm. (Kesten-Spitzer-Whitman)

Also want law of the iterated log. and no points of increase.  
For iter. log., may want Azuma inequality.

Also arcsine laws for last visit to 0 and for positive time fraction.

# Lecture 1 (21.2.11)

Course will focus on Properties of Random Walk & BM.

In first half, discuss Random Walks focusing on SRW in  $\mathbb{Z}$  and  $\mathbb{Z}^d$ .

In second half, define and investigate BM.

Scribe to every lecture. Possibly some HW returned at

end of semester, possibly presentations by students.

Ask if know martingales.

Random Walks:  $S_n = X_1 + \dots + X_n$  for IID  $X_n$ .

$$\mathbb{P} = \mathbb{P}^0 \\ \mathbb{E} = \mathbb{E}^0$$

In most of our examples  $X_i \in \mathbb{R}^d$  or  $\mathbb{Z}^d$ . Sometimes more general. Sometimes talk in generality of Markov chains on

a finite or countable state space.  $\mathbb{P}^x, \mathbb{E}^x$  measures when starting at  $x$ .

Simple random walk (SRW):  $X_i \in \{-1, 1\}$  with equal prob.

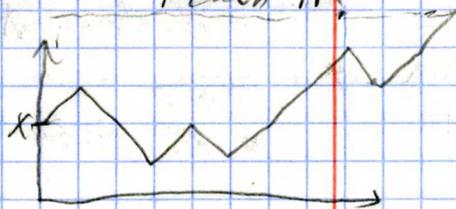
We will start by investigating basic features of SRW.

Gambler's ruin: Gambler has  $x$  shillings, bet fairly one

at a time until reach 0 or  $A$ . What is the chance to reach  $A$ ?

More precisely:  $S_n$  SRW,  $T = \min\{n \mid S_n \in \{0, A\}\}$ .

Calculate  $\mathbb{P}^x(S_T = A)$ .



Remark:  $\mathbb{P}(T < \infty) = 1$  since run of  $A+1$  will occur eventually.

Define  $u(x) = \mathbb{P}^x(S_T = A)$ . Then  $u(0) = 0, u(A) = 1$ ,

$$u(x) = \frac{1}{2}(u(x+1) + u(x-1)) \quad 0 < x < A$$

Many ways to solve, e.g.  $u(x+1) - u(x) = u(x) - u(x-1)$

$$\Rightarrow u(x) = ax + b \Rightarrow u(x) = \frac{x}{A}$$

Remark:  $u$  is a harmonic function (on  $\mathbb{Z}$ ). This is the tip of the iceberg in the relation between RW and certain PDE. Probabilistic Potential theory.

2) What is the expected duration of the game,  $\mathbb{E}^x T$ ?

Define  $v(x) = \mathbb{E}^x T$ , then  $v(0) = 0, v(A) = 0, v(x) = 1 + \frac{1}{2}(v(x+1) + v(x-1))$

$$\mathbb{E}^x T = \mathbb{E}^x \mathbb{E}^x(T \mid X_1) = \mathbb{E}^x(1 + v(x+X_1)) = 1 + \frac{1}{2}(v(x+1) + v(x-1))$$

$$\text{or } v(x+1) + v(x-1) - 2v(x) = -2$$

General Sol.:

$V(x) = ax + b + f(x)$  where  $f(x)$  is any fcn. such that  $f(x+1) + f(x-1) - 2f(x) = -2$ .

GUESS  $f(x) = -x^2$  (operator is like second derivative).

$$\text{Get } V(x) = ax + b - x^2, \quad V(0) = b = 0 \\ V(A) = aA + b - A^2 = 0 \Rightarrow \begin{matrix} b=0 \\ a=A \end{matrix} \Rightarrow V(x) = x(A-x)$$

Exercise: Solve gambler's ruin for  $p < q$  biased walk using matrix diagonalization.

Remark: starting at  $\frac{A}{2}$ ,  $ET$  is quadratic in  $A$  (maybe give as motivating question to class)

Recurrence of SRW: We obtained  $P^x(\text{reach } 0) = 1$  (by taking  $A \rightarrow \infty$ )  
and reflection invariance.

Translation invariance implies  $P^x(\text{reach } y) = 1$ .

Let  $\nu_y = \nu(y) = \min \{n \mid S_n = y\}$  be the first passage time to  $y$ .

It is intuitive that the walk restarts at  $\nu_y$  and hence

(recurrence)  $P^x(\text{reach } y \text{ infinitely often}) = 1$ .

For this we need:

(general) (or even Markov chain)

DEF.: A stopping time for a random walk is a random

variable  $\tau$  taking values in  $\{0, 1, 2, \dots\} \cup \{\infty\}$  such that

for every (finite)  $n$ ,  $\{\tau \leq n\}$  is a measurable fcn. of  $S_1, \dots, S_n$ .

(in other words, if  $\mathcal{F}_n = \sigma(S_1, \dots, S_n)$  then  $\{\tau \leq n\} \in \mathcal{F}_n$ )  
ALSO get  $\{\tau \leq n\} \in \mathcal{F}_n, \{\tau > n\} \in \mathcal{F}_n$

EX.:  $\tau \wedge n$  is a stopping time when  $\tau$  and  $n$  are.  $n$  is a stop. time.  
Hence  $\tau \wedge n$  is a stop. time.

Prop. (Strong Markov property): If  $\tau$  is a stopping time

for  $\{S_n\}_{n \geq 0}$  then  $\{S_{\tau+k} - S_\tau\}$  is also a random walk with the same

step dist., started at 0 and ind. of  $\mathcal{F}_{\tau}$ .

more precisely,  $\forall k \in \mathbb{N}$ ,  $S_1, \dots, S_k, S_1^*, \dots, S_k^*$  we have

$$P^x(\mathcal{E}_{S_1=S_1, \dots, S_k=S_k} \cap \{\tau=k\} \cap \mathcal{E}_{S_{\tau+1}-S_\tau=S_1^*, \dots, S_{\tau+k}-S_\tau=S_k^*}) = P^0(\mathcal{E}_{S_1=S_1^*, \dots, S_k=S_k^*})$$

$$P^x(\mathcal{E}_{S_1=S_1, \dots, S_k=S_k} \cap \{\tau=k\}) P^0(\mathcal{E}_{S_1=S_1^*, \dots, S_k=S_k^*})$$

Remark: We can define the  $\sigma$ -field of events known up to  $\tau$ ,  $\mathcal{F}_\tau$ ,

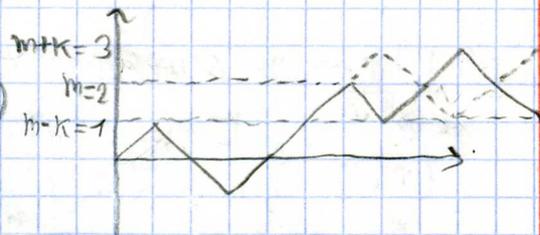
as all  $A \in \mathcal{F}$  such that  $A \cap \{\tau \leq n\} \in \mathcal{F}_n$ . Then  $\forall A \in \mathcal{F}_\tau$ ,

$$P^x(A \cap \{\tau < \infty\} \cap \mathcal{E}_{S_{\tau+1}-S_\tau=S_1^*, \dots, S_{\tau+k}-S_\tau=S_k^*}) = P^x(A \cap \{\tau < \infty\}) P^0(\mathcal{E}_{S_1=S_1^*, \dots, S_k=S_k^*})$$

Above recurrence is an immed. cor.



2) Maximum: Let  $M_n = \max_{0 \leq j \leq n} S_j$   
 Prop: (Dist. of max. and hitting times)  
 For  $k \geq 0$ ,  $P(M_n \geq m, S_n = m-k) = P(S_n = m+k)$



Since if we let  $\tau(m) = \min\{j : S_j = m\}$

and define  $S_j^* = \begin{cases} S_j & j \leq \tau(m) \\ 2m - S_j & j \geq \tau(m) \end{cases}$

then  $S_j^*$  is again a SRW (Andre's Refl. Princ.) by strong Markov prop.

In particular:

Prop: (Dist. of maximum)  $P(M_n \geq m) = P(\tau(m) \leq n) = P(S_n = m) + 2P(S_n > m) \leq 2P(S_n \geq m)$

Proof:  $\sum_{k=-\infty}^{\infty} P(M_n \geq m, S_n = m-k) = \sum_{k=0}^{\infty} P(S_n = m+k) + \sum_{k=1}^{\infty} P(S_n = m+k)$

Exercise: derive  $P(\tau(m) = 2n-1)$

Wald identities: Assume  $S_n$  is a general 1D RW with  $S_0 = 0$ .

First identity: If  $E|X_1| < \infty$ ,  $\mu = EX_1$ , then  $\forall$  stop. time  $\tau$  with  $E\tau < \infty$  we have  $ES_\tau = \mu E\tau$ .

Second identity: If  $E|X_1|^2 < \infty$ ,  $\mu = EX_1$ ,  $\sigma^2 = \text{Var}(X_1)$  then  $\forall$  stop. time  $\tau$  with  $E\tau < \infty$  we have  $E(S_\tau - \mu\tau)^2 = \sigma^2 E\tau$ .

Third identity: Assume that the mgf  $\varphi(\theta) = Ee^{\theta X_1}$  is finite at  $\theta$ , then for every bdd. stop. time  $\tau$ ,  $E\left(\frac{e^{\theta S_\tau}}{\varphi(\theta)^\tau}\right) = 1$

Remark:  $E\tau < \infty$  not suff. for third identity.

If  $E\tau = \infty$ , first identity may fail. E.g.  $\tau =$  hitting time of 1 for SRW.

$1 = ES_\tau \neq 0 = \mu \cdot E\tau = 0 \cdot \infty$ . In fact, for SRW and first identity, suff. that  $E\tau^{1/2} < \infty$ .

Example: Gambler's ruin:  $S_n$  SRW.  $T =$  hitting time of  $\{0, A\}$ .

2 pp. 9.5  
 1/25/30 pp  
 7.5/16 14.5  
 1 pp  
 25.4  
 16.4  
 Lecture 2 (28.2.2010)

$E^x(S_T - x) = \mu E\tau = 0$  (easy to see  $E\tau < \infty$ ).

$\Rightarrow (A-x)P(S_T = A) + (-x)P(S_T = 0) = 0 \Rightarrow P(S_T = A) = \frac{x}{A}$

$E^x(S_T - x)^2 = \sigma^2 E\tau = E\tau \Rightarrow E\tau = \frac{x}{A}(A-x)^2 + \frac{A-x}{A}(-x)^2 = x(A-x)$

Proofs of 1st and 3rd identity (and like 1st): (special cases of mart. thms.)

1st identity: If  $\tau$  is bdd. by  $M$ :  $S_\tau = \sum_{n=0}^M S_n \mathbb{1}_{\tau \geq n} = \sum_{n=1}^M X_n \mathbb{1}_{\tau \geq n}$   
 $ES_\tau = \sum_{n=0}^M E(X_n \mathbb{1}_{\tau \geq n}) = \sum_{n=1}^M E(X_n) P(\tau \geq n) = \mu \sum_{n=1}^M P(\tau \geq n) = \mu E\tau$

(since  $\forall$  non-neg. RV  $Y$ ,  $EY = \sum_{n=1}^{\infty} P(Y \geq n)$ )

If  $\tau$  not bdd., need to exchange  $E$  and sum. This is possible since

$\sum_{n=1}^{\infty} E|X_1| \mathbb{1}_{\tau \geq n} = \sum_{n=1}^{\infty} E|X_1| \cdot P(\tau \geq n) = E|X_1| \cdot E\tau < \infty$

Ask about make-up lecture  
 114 (Lecture 2)  
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 (Lecture 50)

3rd identity: suppose  $\tau \leq M$  w/prob. 1.

$$\begin{aligned} \mathbb{E}\left[\frac{e^{\theta S_\tau}}{\varphi(\theta)^\tau}\right] &= \sum_{n=0}^M \mathbb{E}\left[\frac{e^{\theta S_\tau}}{\varphi(\theta)^\tau} \mathbb{1}_{\{\tau=n\}}\right] = \sum_{n=0}^M \mathbb{E}\left[\frac{e^{\theta S_n}}{\varphi(\theta)^n} \mathbb{1}_{\{\tau=n\}}\right] = \\ &= \sum_{n=0}^M \mathbb{E}\left[\frac{e^{\theta S_n}}{\varphi(\theta)^n} \mathbb{1}_{\{\tau=n\}}\right] \cdot \mathbb{E}\left[\frac{e^{\theta S_{M-n}}}{\varphi(\theta)^{M-n}}\right] = \sum_{n=0}^M \mathbb{E}\left[\frac{e^{\theta S_M}}{\varphi(\theta)^M} \mathbb{1}_{\{\tau=n\}}\right] = \mathbb{E}\left[\frac{e^{\theta S_M}}{\varphi(\theta)^M}\right] = 1_0 \end{aligned}$$

$\tau$  is a stop. time

Derivation of first passage time mgf using 3rd identity

Let  $\tau = \min\{n \geq 0 \mid S_n = 1\}$  for a SRW  $(S_n)_{n \geq 0}$ .

We would like to calculate the mgf  $\mathbb{E}(S^\tau)$ .

If  $\tau$  was bdd., we could write  $\mathbb{E}\left[\frac{e^{\theta S_\tau}}{\varphi(\theta)^\tau}\right] = 1$  where  $\varphi(\theta) = \mathbb{E}[e^{\theta X_1}]$ .

First, let us see that this is sufficient to find  $\mathbb{E}(S^\tau)$ .

By def.,  $S_\tau = 1$ , thus we have  $\mathbb{E}\left[\frac{1}{\varphi(\theta)^\tau}\right] = e^{-\theta}$ .

Note that  $\varphi(\theta) = \frac{1}{2}(e^\theta + e^{-\theta}) = \cosh(\theta)$ . Writing  $s = \frac{1}{\varphi(\theta)} = \frac{2}{e^\theta + e^{-\theta}}$

We have  $s(e^\theta + e^{-\theta}) = 2 \Rightarrow s(1 + e^{-2\theta}) = 2e^{-\theta} \Rightarrow e^{-2\theta}s - 2e^{-\theta} + s = 0$

$$e^{-\theta} = \frac{2 \pm \sqrt{4 - 4s^2}}{2s} = \frac{1 \pm \sqrt{1 - s^2}}{s}. \text{ The two sols. correspond to } e^\theta \text{ and } e^{-\theta}$$

Restricting to  $\theta > 0$  we can write  $e^{-\theta} = \frac{1 - \sqrt{1 - s^2}}{s}$  as the smaller sol. (since  $s \in (0, 1)$ )

Thus  $\mathbb{E}(S^\tau) = \frac{1 - \sqrt{1 - s^2}}{s}$  for  $s \in (0, 1)$  (and thus for  $|s| < 1$  by analytic cont.)

It remains to show  $\mathbb{E}\left[\frac{e^{\theta S_\tau}}{\varphi(\theta)^\tau}\right] = 1$  for  $\theta > 0$ . By the 3rd Wald

identity  $\mathbb{E}\left[\frac{e^{\theta S_{\tau \wedge n}}}{\varphi(\theta)^{\tau \wedge n}}\right] = 1 \quad \forall n \geq 0$ , since  $\mathbb{P}(\tau < \infty) = 1$ ,

$$\lim_{n \rightarrow \infty} \frac{e^{\theta S_{\tau \wedge n}}}{\varphi(\theta)^{\tau \wedge n}} = \frac{e^{\theta S_\tau}}{\varphi(\theta)^\tau} \text{ a.s. . Noting that } 0 < \frac{1}{\varphi(\theta)} \leq 1 \text{ and}$$

$0 < e^{\theta S_{\tau \wedge n}} \leq e^\theta$ , the equality follows by the dominated conv. thm.

Sample path prop. of SRW

Lemma:  $\mathbb{P}(S_1 > 0, \dots, S_{2n} > 0) = \mathbb{P}(S_{2n} = 0)$

Proof: Enough to show

$$\mathbb{P}(S_1 > 0, \dots, S_{2n} > 0) = \frac{1}{2} \mathbb{P}(S_{2n} = 0) \quad (r \geq 1)$$

Already saw in ballot theorem:  $\mathbb{P}(S_1 > 0, \dots, S_{2n-1} > 0, S_{2n} = 2r) = \frac{r}{n} \mathbb{P}(S_{2n} = 2r) =$

$$= \frac{1}{2} (\mathbb{P}(S_{2n-1} = 2r-1) - \mathbb{P}(S_{2n-1} = 2r+1))$$

part of ballot thm. proof

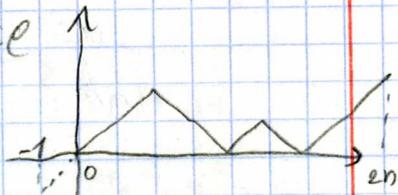
Thus:  $P(S_1 > 0, \dots, S_{2n} > 0) = \sum_{r=1}^n \frac{1}{2} (P(S_{2n-1} = 2r-1) - P(S_{2n-1} = 2r+1)) = \frac{1}{2} P(S_{2n-1} = 1)$

and we are done since  $P(S_{2n} = 0) = \frac{1}{2} (P(S_{2n-1} = 1) + P(S_{2n-1} = -1)) = P(S_{2n-1} = 1)$ .

Note also that  $P(S_1 \geq 0, \dots, S_{2n} \geq 0) = P(S_{2n} = 0)$  since

$\frac{1}{2} P(S_{2n} = 0) = P(S_1 > 0, \dots, S_{2n} > 0) = \frac{1}{2} P(S_1 \geq 0, \dots, S_{2n-1} \geq 0) =$

Parity  $\rightarrow \frac{1}{2} P(S_1 \geq 0, \dots, S_{2n} \geq 0)$



This gives another version of the first passage time dist.

Let  $T = \min(n \geq 1 | S_n = 0)$  then  $P(T = 2n) = P(S_1 \neq 0, \dots, S_{2n-2} \neq 0) - P(S_1 \neq 0, \dots, S_{2n} = 0)$

$= P(S_{2n-2} = 0) - P(S_{2n} = 0) = \frac{1}{2n-1} P(S_{2n} = 0)$ .

Arcsine laws - last zero and time above the axis

Consider last visit to zero up to time  $2n$ . Surprisingly, dist.

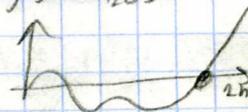
is sym. around  $n$  and achieves max. at edges.

For example (Feller) if tossing a coin every second for a year,

$P(\text{last } 0 \text{ before } 9 \text{ days}) \approx \frac{1}{10}, P(\text{before } 2\frac{1}{7} \text{ days}) \approx \frac{1}{20}$ ,

$P(\text{before } 2 \text{ hours and } 10 \text{ minutes}) \approx \frac{1}{100}$ .

(arcsine dist. for last zero)



Thm: Let  $T = \max(k \leq n | S_{2k} = 0)$  (T is not a stop. time!)

then  $P(T = 2k) = P(S_{2k} = 0) P(S_{2n-2k} = 0)$ .

Remark:  $P(S_{2k} = 0) = \binom{2k}{k} 2^{-2k} \sim \frac{1}{\sqrt{\pi k}}$  as  $k \rightarrow \infty$ .

thus  $P(T = 2k) \approx \frac{1}{\pi \sqrt{k(n-k)}} = \frac{1}{n} \cdot \frac{1}{\pi \sqrt{\frac{k}{n}(1-\frac{k}{n})}}$ . Note that

$\int_0^x \frac{1}{\pi \sqrt{y(1-y)}} dy = \frac{2}{\pi} \arcsin(\sqrt{x})$

Corollary:  $\lim_{n \rightarrow \infty} \lim_{k/n \rightarrow x} P(T \leq 2k) = \frac{2}{\pi} \arcsin(\sqrt{x})$ .

Markov Prop.

Proof of thm:  $P(T = 2k) = P(S_{2k} = 0, S_{2k+1} \neq 0, \dots, S_{2n} \neq 0) =$

$= P(S_{2k} = 0) \cdot P(S_1 \neq 0, \dots, S_{2n-2k} \neq 0) \stackrel{\text{lemma}}{=} P(S_{2k} = 0) \cdot P(S_{2n-2k} = 0)$

Thm. (arcsine dist. for time above axis):  $\{0 \leq m \leq 2n-1 | S_m \geq 0, S_{m+1} \geq 0\}$

$P(\text{time spent above axis} = 2k) = P(S_{2k} = 0) P(S_{2n-2k} = 0)$ .



similar Corollary

Example (Feller): as before, if 2 players play every second for a year,

there is chance  $\frac{1}{10}$  that the less fortunate player will lead for less than 9 days.

Denote  $P(\text{time above axis} = 2k) =: b_{2k, 2n}$   
Proof of thm. For  $k=n$   $b_{2n, 2n} = P(S'_1 \geq 0, \dots, S'_{2n} \geq 0) = P(S'_{2n} = 0)$   
 same for  $k=0$ . Now take  $0 < k < n-1$ . Denote also

$f_{2r} = P(\text{first return to 0 at time } 2r)$ . Then

$$b_{2k, 2n} = P(\text{time above axis} = 2k) = \sum_{r=1}^n P(\text{time} = 2k, \text{first return} = 2r) =$$

$$= \sum_{r=1}^n f_{2r} P(\text{time} = 2k \mid \text{first return} = 2r) \stackrel{\text{symmetry}}{=} \frac{1}{2} \sum_{r=1}^k f_{2r} b_{2k-2r, 2n-2r} + \frac{1}{2} \sum_{r=1}^{n-k} f_{2r} b_{2k, 2n-2r}$$

Assume by induction that for all  $1 \leq \nu \leq n-1$  we have

shown  $b_{2k, 2\nu} = P(S'_{2k} = 0) P(S'_{2\nu-2k} = 0)$ . Then

$$b_{2k, 2n} = \frac{1}{2} \sum_{r=1}^k f_{2r} (P(S'_{2k-2r} = 0) P(S'_{2n-2k} = 0)) + \frac{1}{2} \sum_{r=1}^{n-k} f_{2r} (P(S'_{2k} = 0) P(S'_{2n-2k-2r} = 0))$$

$$= \frac{1}{2} P(S'_{2n-2k} = 0) \sum_{r=1}^k f_{2r} P(S'_{2k-2r} = 0) + \frac{1}{2} P(S'_{2k} = 0) \sum_{r=1}^{n-k} f_{2r} P(S'_{2n-2k-2r} = 0) =$$

$$= \frac{1}{2} P(S'_{2n-2k} = 0) P(S'_{2k} = 0) + \frac{1}{2} P(S'_{2k} = 0) P(S'_{2n-2k} = 0)$$

Time reversal

There are 2 more arc-sine laws obtained from the first two using time reversal.

Denoting the reversed path by  $\{S_j^*\}$  we have  $S_j^* = S_n - S_{n-k}$

Thus, for example,  $\{S_j^* > 0 \forall 1 \leq j \leq n\} = \{S_n > S_j \forall 0 \leq j \leq n-1\}$  and

Weak ladder point  
 Also  $P(S_{2n} \geq s_j) \stackrel{V.3}{=} P(S'_{2n} = 0)$   
 similarly

we have  $P(S'_{2n} > s_j \forall 0 \leq j \leq 2n-1) = P(S_j^* > 0 \forall 1 \leq j \leq 2n) = \frac{1}{2} P(S_{2n} = 0)$

at time  $2n$  (and a similar formula for records at odd times)

arc-sine law for first visit to terminal point:

$$P(\min(j \mid S_{2j} = S_{2n}) = k) = P(S'_{2k} = 0) P(S'_{2n-2k} = 0)$$

dual to dist. of last zero.

arc-sine law for position of first/last global maximum

For simplicity, take an even length walk and suppose the <sup>first</sup> maximum occurs at  $2k$ . Then we have:

$$\underbrace{S_0 < S_{2k}}_{\text{Prob.} = \frac{1}{2} P(S_{2k} = 0)}, \underbrace{S_{2k-1} < S_{2k}}_{\text{Prob.} = \frac{1}{2} P(S_{2k} = 0)}, \underbrace{S_{2k+1} < S_{2k}}_{\text{Prob.} = \frac{1}{2} P(S_{2k} = 0)}, \dots, \underbrace{S_{2n} < S_{2k}}_{\text{Prob.} = \frac{1}{2} P(S_{2k} = 0)}$$

Thus  $P(\text{first max. at } 2k) = \frac{1}{2} P(S_{2k} = 0) P(S_{2n-2k} = 0)$

The same is true for last max. at  $2k$  (moving the  $\frac{1}{2}$  between the factors).  
 A similar statement holds for max. at an odd location.

Random walk bridge equidistribution of positive time

In contrast to the arc-sine law for ordinary SRW, for

d, SRW bridge we have (an analogous thm. holds also for the pos. of Thm.:  $P(\text{time spent above axis} = 2k | S'_{2n} = 0) = \frac{1}{n+1}$  by considering global maximum

Let  $C_{2k, 2n} = P(\text{time above axis} = 2k | S'_{2n} = 0) = P(\text{time above axis} = 2k, S'_{2n} = 0)$

For  $k=n$ , by ballot thm.  $J_{2n, 2n} = \frac{1}{2n+1} \binom{2n+1}{n+1} 2^{-2n} = \frac{1}{n+1} \binom{2n}{n} 2^{-2n} = \frac{1}{n+1} P(S'_{2n} = 0)$

Similarly for  $k=0$ , let  $1 \leq k < n$ ,  $F_{2r} = P(\text{first return to 0 is at } 2r)$

$J_{2k, 2n} = \sum_{r=1}^n P(\text{time} = 2k, S'_{2n} = 0, \text{first return} = 2r) = \frac{1}{2} \sum_{r=1}^k F_{2r} J_{2k-2r, 2n-2r} + \frac{1}{2} \sum_{r=1}^{n-k} F_{2r} J_{2k, 2n-2r}$

By induction:  $J_{2k, 2n} = \frac{1}{2} \sum_{r=1}^k F_{2r} \cdot \frac{P(S'_{2n-2r} = 0)}{n-r+1} + \frac{1}{2} \sum_{r=1}^{n-k} F_{2r} \cdot \frac{P(S'_{2n-2r} = 0)}{n-r+1}$

Recalling that  $F_{2r} = \frac{1}{2} \cdot \frac{1}{r} P(S'_{2r-2} = 0)$  by ballot thm. cor. again, we have

$J_{2k, 2n} = \frac{1}{4} \sum_{r=1}^k \frac{1}{r(n-r+1)} P(S'_{2r-2} = 0) P(S'_{2n-2r} = 0) + \frac{1}{4} \sum_{r=1}^{n-k} \frac{P(S'_{2r-2} = 0) P(S'_{2n-2r} = 0)}{r(n-r+1)}$

replace  $n$  by  $n-r+1$  in second sum  $\rightarrow = \frac{1}{4} \sum_{r=1}^n \frac{P(S'_{2r-2} = 0) P(S'_{2n-2r} = 0)}{r(n-r+1)}$  ind. of  $k$ !

Lecture 3 (7.3.11)

name, email, registered, may register

Reminders: no class next week. Pass contact page. Find a scribe.

Mention that results of E. Sørensen Andersen give a combinatorial approach to proving some of the arcsine laws for general  $z$ -RW even without moment assumptions.

Law of the iterated logarithm

For a SRW  $(S_n)$ , by the CLT, we have  $S_n$  is approx.  $N(0, n)$ .

Since if  $N \sim N(0, n)$  we have  $P(N \in (k - \frac{1}{2}, k + \frac{1}{2})) \approx \frac{1}{\sqrt{2\pi n}} e^{-k^2/2n}$

We may expect the same for  $S_n$ . Note however that  $S_n$  necessarily

has the parity of  $n$ . Thus, we may expect  $P(S_n = k) \propto \begin{cases} 0 & k \text{ of odd parity from } n \\ \frac{1}{\sqrt{2\pi n}} e^{-k^2/2n} & \text{O/W} \end{cases}$

Similarly,  $P(N > k) \sim \frac{n}{k} \cdot \frac{1}{\sqrt{2\pi n}} e^{-k^2/2n}$  and we may expect something similar from  $S_n$ .

(LCLT)

Lemma: i) For  $k = o(n^{3/4})$  of the same parity as  $n$ ,  $P(S_n = k) \sim \frac{1}{\sqrt{2\pi n}} e^{-k^2/2n}$  ( $n \rightarrow \infty$ )

ii) For all  $n, k \geq 0$   $P(S_n \geq k) \leq e^{-k^2/2n}$

Proof: i)  $P(S_n = k) = \binom{n}{(n+k)/2} 2^{-n} = \frac{n! 2^{-n}}{\frac{(n+k)!}{2^{(n+k)/2}} \frac{(n-k)!}{2^{(n-k)/2}}} = \frac{n! 2^{-n}}{\frac{(n+k)! (n-k)!}{2^n}} = \frac{n! 2^{-n}}{(n+k)! (n-k)!} = \frac{1}{\sqrt{2\pi n}} e^{-k^2/2n} \cdot \frac{n!}{(n+k)! (n-k)!} \cdot \sqrt{2\pi n} e^{k^2/2n}$

$k$  and  $n$  have same parity

$\exp(n \log n - \frac{n+k}{2} \log(n+k) - \frac{n-k}{2} \log(n-k))$

$$\log(1-x) = -\left(x + \frac{x^2}{2} + \frac{x^3}{3} + O(x^4)\right)$$

$$\log(1+\frac{k}{n}) - \log(1-\frac{k}{n}) = -\frac{k}{n} + \frac{k^2}{2n^2} - \frac{k^3}{3n^3} + O(\frac{k^4}{n^4}) + \frac{k}{n} + \frac{k^2}{2n^2} + \frac{k^3}{3n^3} + O(\frac{k^4}{n^4}) = \frac{k^2}{n} + O(\frac{k^4}{n^3})$$

Thus  $P(S_n = k) \sim \sqrt{\frac{2}{\pi n}} e^{-\frac{k^2}{2n}} + O(\frac{k^4}{n^3}) \sim \sqrt{\frac{2}{\pi n}} e^{-\frac{k^2}{2n}}$

ii) special case of Bernstein-Chernoff-Hoeffding-Azuma inequality.

Note  $\frac{1}{2}e^\theta + \frac{1}{2}e^{-\theta} \leq e^{\theta^2/2}$  by comparing Taylor coef. ( $\frac{\theta^{2n}}{(2n)!} \leq \frac{\theta^{2n}}{2^n n!}$ )

$$P(S_n \geq k) \leq \frac{E e^{\theta S_n}}{e^{\theta k}} = \frac{(\frac{1}{2}e^\theta + \frac{1}{2}e^{-\theta})^n}{e^{\theta k}} \leq e^{\theta^2 n/2 - \theta k} = e^{-k^2/2n}$$

$\theta = \frac{k}{n} \quad (k \geq 0)$

Reminder:  $M_n = \max_{0 \leq j \leq n} S_j$ ,  $P(M_n \geq k) = P(S_n = k) + 2P(S_n > k)$

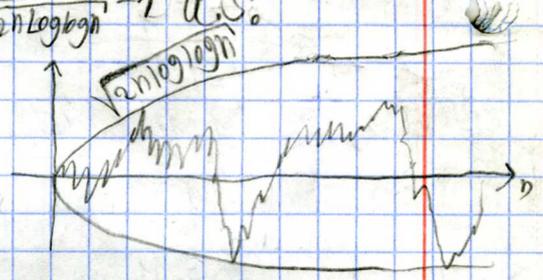
Exercise  $P(A_n \text{ occur inf. often}) \geq \limsup P(A_n)$ . Gives lower bound of  $P(M_n)$  in LIL.

Theorem (Law of the iter. lg.)  $\limsup_{n \rightarrow \infty} \frac{S_n}{\sqrt{2n \log \log n}} = 1$  a.s.

Remark: By symmetry  $\liminf_{n \rightarrow \infty} \frac{S_n}{\sqrt{2n \log \log n}} = -1$  a.s.

We will ignore issues of integrality of  $n$  in  $S_n$  in the proof.

Idea:  $P(S_{a^k} \geq \sqrt{2a^k \log \log(a^k)}) \approx e^{-\frac{1}{k \log a}}$



$\left(\frac{1}{k \log a}\right)^2$  summable iff  $a > 1$ .

These are relatively inf. events and the walk does not fluctuate much between the  $a^k$ . This already gives upper bound on subsequence.

Lemma: For any  $\epsilon \in (0, 1)$ ,  $\limsup_{n \rightarrow \infty} \max_{n \leq j \leq (1+\epsilon)n} \frac{|S_j - S_n|}{\sqrt{2n \log \log n}} \leq 4\sqrt{\epsilon}$  a.s.

Proof: Fix  $\epsilon, \alpha > 0$ ,  $n_k = (1+\epsilon)^k$ , then

$$P\left(\max_{n_k \leq j \leq n_{k+1}} |S_j - S_{n_k}| \geq \alpha \sqrt{2n_k \log \log(n_k)}\right) \leq 4P\left(S_{n_{k+1} - n_k} \geq \alpha \sqrt{2n_k \log \log(n_k)}\right) \leq 4e^{-\frac{\alpha^2}{\epsilon}}$$

Use better proof below!

So if  $\alpha^2 > \epsilon$ , only finitely many of these events occur by the Borel-Cantelli lemma. Thus for all large  $k$

$$\max_{n_k \leq j \leq n_{k+1}} |S_j - S_{n_k}| \leq \alpha U(n_k) \quad (\alpha^2 > \epsilon)$$

where  $U(t) = \sqrt{2t \log \log t}$

Noting that for any  $q > 1$ ,  $\frac{U(q^{k+1})}{q^{k+1}} = \frac{U(q^k)}{q^k} \cdot \frac{q^k}{q^{k+1}} \leq \frac{1}{q}$  and  $k$  large enough.

Since  $\frac{U(t)}{t}$  is decreasing in  $t$  for  $t$  large enough.

We see that if  $n_k \leq n \leq n_{k+1}$  and  $n \leq j \leq (1+\epsilon)n$  and  $k$  is large:

$$|S_j - S_n| \leq |S_j - S_{n_{k+1}}| + |S_{n_{k+1}} - S_{n_k}| + |S_{n_k} - S_n| \leq 2\alpha U(n_k) + \alpha U(n_{k+1}) \leq$$

$$\left( \leq (2\alpha + \alpha(1+\epsilon))U(n_k) \leq 4\alpha U(n_k) \right)$$

Taking  $\alpha \downarrow \sqrt{\epsilon}$  over a sequence proves the lemma.

Proof of thm: For any  $\delta > 0$ ,  $\limsup_{a \rightarrow \infty} S_{a^k} \leq (1+\delta)U(a^k)$  by the proof idea above and Borel-Cantelli. Taking  $a = 1/\epsilon$  we have by the lemma

$$\text{for } a^k \leq n \leq a^{k+1}: \frac{S_n}{U(n)} = \frac{S_{a^k}}{U(a^k)} \frac{U(a^k)}{U(n)} + \frac{S_n - S_{a^k}}{U(a^k)} \frac{U(a^k)}{U(n)} \leq 1 + \delta + 4\sqrt{\epsilon}$$

Taking  $\epsilon, \delta \rightarrow 0$  over a subsequence proves the upper bound in the thm.

For the lower bound note that  $P(S_{a^k} - S_{a^{k-1}} \geq (1-\delta)U(a^k - a^{k-1})) \geq$

$$\geq \sqrt{\frac{2}{\pi(a^k - a^{k-1})}} e^{-\frac{(1-\delta)^2 \log \log(a^k - a^{k-1})}{a.s.}} \text{ which is not summable in } k. \text{ Hence}$$

by the B-G lemma, for large  $k$ ,  $S_{a^k} - S_{a^{k-1}} \geq (1-\delta)U(a^k - a^{k-1})$ . H

By the upper bound for large  $k$ ,  $S_{a^k} \leq (1+\epsilon)U(a^k)$ . Thus

$$\frac{S_{a^k}}{U(a^k)} \geq (1-\delta) \frac{U(a^k - a^{k-1})}{U(a^k)} + \frac{S_{a^{k-1}}}{U(a^k)} \geq (1-\delta) \frac{U(a^k - a^{k-1})}{U(a^k)} - (1+\epsilon) \frac{U(a^{k-1})}{U(a^k)} \xrightarrow{k \rightarrow \infty} \frac{(1-\delta)\sqrt{1-\frac{1}{a}} - (1+\epsilon)}{\sqrt{a}}$$

Thus  $\limsup_{n \rightarrow \infty} \frac{S_n}{U(n)} \geq (1-\delta)\sqrt{1-\frac{1}{a}} - \frac{1+\epsilon}{\sqrt{a}}$ . Taking  $\delta \rightarrow 0, a \rightarrow \infty$  over

a subsequence proves the lower bound in the thm.

Better proof for upper bound:  $P(\max_{0 \leq n \leq a^k} S_n \geq (1+\delta)U(a^k)) \leq 2e^{-\frac{(1+\delta)^2 U(a^k)^2}{2a^k}} = 2 \left( \frac{1}{k \log a} \right)^{(1+\delta)^2}$  is summable, so occurs only finitely many times a.s.

For large  $n$ , write  $a^{k-1} \leq n \leq a^k$  and note

$$\frac{S_n}{U(n)} = \frac{S_n}{U(a^k)} \cdot \frac{U(a^k)}{a^k} \cdot \frac{a^k}{n} \leq (1+\delta) \cdot a \quad \text{since } \frac{U(n)}{n} \geq \frac{U(a^k)}{a^k} \text{ for large } n.$$

Thus  $\limsup_{n \rightarrow \infty} \frac{S_n}{U(n)} \leq (1+\delta)a$  a.s. Taking  $\delta \rightarrow 0, a \rightarrow \infty$  over a

sequence, the upper bound is proved.

Higher dimensions and more general RW (Durrett Chap 4) otherwise transient

Recurrence/transience: A RW taking values in  $\mathbb{R}^d$  is called recurrent if  $P(S_n = 0 \text{ i.o.}) = 1$ .

We also define the set  $P_i = \{x \in \mathbb{Z}^d \mid \exists n \text{ s.t. } P(S_n = x) > 0\}$ , the set of possible values for  $S_n$ . A state  $x$  is called recurrent if  $P(S_n = x \text{ i.o.}) = 1$ .

The strong Markov property implies that the walk is recurrent iff all its possible states are recurrent.

EX.: If one state is recurrent then all possible states are recurrent.

Can it be that  $P(S_n = 0 \text{ i.o.}) < 1$ ? The next thm. shows it is impossible.

Put discussion after next page.

say that an event is permutable if its occurrence for a walk with increments  $(X_1, X_2, \dots)$  is unaffected by applying a finite perm. to the increments. In other words, if our increments take values in  $S$  (in our case,  $\mathbb{Z}^d$  or  $\mathbb{R}^d$ ) and we take  $\mu = S^{\mathbb{N}}$  with the prod. measure coming from our increments then  $A \in \mathcal{F}$  is perm. if  $\pi A = \{\omega \mid \pi^{-1} \omega \in A\} = A$  for all finite perm.  $\pi$ .

The collection of permutable events forms the exchangeable  $\sigma$ -field  $\mathcal{E} \subset \mathcal{F}$ . For example:  $A = \{S_n = 0 \text{ i.o.}\}$ ,  $A = \{\limsup \frac{S_n}{\sqrt{n}} \geq 1\}$ . All tail events are permutable so tail  $\sigma$ -field  $\subset \mathcal{E}$ .

Thm. (Hewitt-Savage 0-1 law): For a RW (or an IID sequence), if  $A \in \mathcal{E}$  then  $P(A) \in \{0, 1\}$ .

Proof: Fix  $A \in \mathcal{E}$ . We will show  $P(A) = P(A)^2$ . Let  $A_n \in \sigma(X_1, \dots, X_n)$

(1) be such that  $P(A_n \Delta A) \rightarrow 0$  (where  $A_n \Delta A = (A \setminus A_n) \cup (A_n \setminus A)$  so that  $1_{A_n \Delta A} = |1_{A_n} - 1_A|$ )

Let  $\pi = \pi_n$  be the perm. which exchanges coords.  $1, \dots, n$  with  $n+1, \dots, 2n$ .

Write  $A'_n = \pi A_n$ . Since  $\pi$  preserves measure

(2)  $P(A_n \Delta A) = P(\pi(A_n \Delta A)) = P(A'_n \Delta A)$  permutable of  $A$ .

Since  $|P(B) - P(C)| \leq P(B \Delta C)$  (triangle inequality in  $1_{B \Delta C} = |1_B - 1_C|$ )

(3) (1) and (2) imply  $P(A_n), P(A'_n) \rightarrow P(A)$ . However, it also implies

$P(A_n \Delta A'_n) \leq P(A_n \Delta A) + P(A'_n \Delta A) \rightarrow 0$ . From which

$0 \leq P(A_n) - P(A_n \cap A'_n) \leq P(A_n \Delta A'_n) - P(A_n \cap A'_n) = P(A_n \Delta A'_n) \rightarrow 0$

implying  $P(A_n \cap A'_n) \xrightarrow{(3)} P(A)$ . Recalling  $P(A_n \cap A'_n) = P(A_n)P(A'_n)$  by ind.,

we are done by (3).

Application: For a RW in  $\mathbb{R}$ , exactly one of the following has Prob. 1:

i)  $S_n = 0 \forall n$ . ii)  $S_n \rightarrow \infty$  iii)  $S_n \rightarrow -\infty$  iv)  $\limsup S_n = \infty$ .

Proof:  $\limsup$  is a constant  $c \in [-\infty, \infty]$  by the prev. thm.

$c \in \{-\infty, \infty\}$  unless  $X_1 = 0$  since  $c = c - X_1$ . The same applies to  $\liminf S_n$ , giving the result.

Ex.: Show that if  $X_1 \in \mathbb{R}$  is non-degenerate and either i) symmetric

ii) Has  $EX_1 = 0, EX_1^2 < \infty$  then we are in case (iv) (actually  $EX_1 = 0$  surprises as we will see)

(Scribe) continuing from recurrence discussion two pages ago.

Fix mistake from last time in LL

Prop.: A RW in  $\mathbb{R}^d$  is point-recurrent iff one of the following holds:

- i)  $P(\exists n \geq 1, S_n = 0) = 1$
- ii)  $P(S_n = 0 \text{ i.o.}) = 1$
- iii)  $\sum_{n=0}^{\infty} P(S_n = 0) = \infty$ .

Proof: By the Strong Markov Property the number of returns to the origin is dist.  $\text{Geom}(P)$  with  $P = P(\text{no return to } 0)$ .

(iii) follows since  $E(\text{number of returns}) = \sum_{n=0}^{\infty} P(S_n = 0) = \frac{1}{P}$ .

Remark: If the RW has a cont. dist., a more useful notion is neighbourhood-recurrence:  $\forall \epsilon > 0, P(|S_n| < \epsilon \text{ i.o.}) = 1$ .

One can similarly define possible values using neighbourhoods and similar Prop. are valid.

Durrett chp 4

(~1920) Polya's thm.: SRW on  $\mathbb{Z}^d$  is recurrent iff  $d=1,2$ . (Durrett cites Kakutani's drunk man/wife

Proof: In 1D we already saw this but we can get this also

by Stirling since  $P(S_n = 0) \sim \frac{1}{\sqrt{\pi n}}$ . For  $d=2$ , note that by rotating the lattice by  $45^\circ$ , the walk is equiv to making one step each turn in each of 2 ind. SRW.



Hence  $P(S_{2n} = 0) = \left(\frac{1}{\sqrt{\pi n}}\right)^2$ . For  $d=3$ :

$$P(S_{2n} = 0) = 6^{-2n} \sum_{\substack{j,k \geq 0 \\ j+k=n}} \frac{(2n)!}{(j!k!(n-j-k)!)} = 2^{-2n} \binom{2n}{n} \sum_{j,k} \left( \frac{3^{-n} n!}{j!k!(n-j-k)!} \right)^2 \leq 2^{-2n} \binom{2n}{n} \max_{j,k} 3^{-n} \frac{n!}{j!k!(n-j-k)!}$$

It remains to show  $\max_{j,k} \dots \leq \frac{C}{n}$  for some  $C > 0$ . Indeed, the max. occurs when the 3 numbers are as close as possible to  $\frac{n}{3}$  and the result follows from Stirling.

For  $d \geq 4$ , observe the walk is transient since we can ignore the last  $d-3$  coord. and it will still be transient.

(~1950) Ghug-Fuchs theorem: A RW in  $\mathbb{R}^d$  is neighbourhood-recurrent

if  $d=1$  and  $\frac{S_n}{n} \rightarrow 0$  in prob. (weak law of large numbers) or if  $d=2$  and  $\frac{S_n}{\sqrt{n}} \Rightarrow$  normal dist.

If  $d \geq 3$  and the random walk is not contained in a plane then it is (neighbourhood-) transient.

Exercise: show in 1D that it does not imply point-rec. For RW in  $\mathbb{R}^d$

(point-recurrent if on  $\mathbb{Z}^d$  follows as cor.)

SLN, KSW thm. and proof of first part on  $\mathbb{Z}^d$  when  $\text{mean} = 0$ . Stable law counterexample

Step (14):  $\exists$  recurrent RW in 1D with arbitrary large tails.

Strong law of large numbers: If  $\mathbb{E}X_1 = \mu$  (in 1D) then  $\frac{S_n}{n} \rightarrow \mu$  a.s.

Implies that if  $\mu \neq 0$  then  $S_n \rightarrow \infty$  or  $S_n \rightarrow -\infty$  a.s. according to sign of  $\mu$ .  
In particular, no recurrence.

Birkhoff ergodic thm. (for factor of IID): Given  $X_1, X_2, \dots$

i.i.d. RVS (say in  $\mathbb{R}$ ) and  $g: \mathbb{R}^2 \rightarrow \mathbb{R}$  measurable.

If we define  $Y_n = g(X_n, X_{n+1}, \dots)$  and if  $\mathbb{E}|Y_1| < \infty$

then  $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n Y_k = \mathbb{E}Y_1$  a.s. and in  $L_1$ .

Remark: sufficient to have the  $Y$ 's be any stat. ergodic seq. with  $\mathbb{E}|Y_1| < \infty$ . We will not prove the thm. in our course. (1967)

Application: Theorem (Kesten-Spitzer-Whitman)

For a RW  $S_n$  on  $\mathbb{R}^d$ , let  $R_n$  be the number of distinct sites visited by time  $n$ ,  $R_n = |\{S_0, \dots, S_n\}|$

then  $\frac{R_n}{n} \rightarrow 0$  (no return to zero) a.s. (and in  $L_1$ ) by bdd. conv.

Remark: true for any Markov chain or even stat. seq.

Proof:  $R_n = \sum_{j=0}^n \mathbb{1}\{S_j \text{ not revisited by time } n\}$ .

$R_n \geq \sum_{j=0}^n \mathbb{1}\{S_j \text{ never revisited}\}$  and by the ergodic thm:

$$\liminf_{n \rightarrow \infty} \frac{R_n}{n} \geq \mathbb{P}(S_0 \text{ never revisited}) = \mathbb{P}(\text{no return to } 0)$$

For the other direction, fix  $M \geq 1$ .

$$R_n \leq \sum_{j=0}^{n-M} \mathbb{1}\{S_j \text{ not revisited by time } j+M\} + M$$

By the ergodic thm:

$$\limsup_{n \rightarrow \infty} \frac{R_n}{n} \leq \mathbb{P}(0 \text{ not revisited by time } M+1) \xrightarrow{M \rightarrow \infty} \mathbb{P}(\text{no return to } 0)$$

since event on right is intersection of events on left.

Proof of recurrence thm. for  $d=1$  when  $\text{mean} = 0$ :

If  $X_1$  has mean 0 then  $\frac{S_n}{n} \rightarrow 0$  a.s. by the SLLN.

Hence for any  $\epsilon > 0$ ,  $\limsup_{n \rightarrow \infty} \frac{R_n}{n} \leq \epsilon$ . Thus  $\frac{R_n}{n} \rightarrow 0 \stackrel{\text{KSW}}{=} \mathbb{P}(\text{no return to } 0) \stackrel{\text{KSW}}{=} 0$ .

Example that SLLN not equiv. to WLLN:  $X_1$ , sym. with  $P(X_1 > x) \sim \frac{1}{x \log x}$  as  $x \rightarrow \infty$

Then weak law holds but strong law does not, can prove by Fourier analysis? However, weak law criterion is not sharp for recurrence.

For cont. RW (and neighbourhood recurrence), can consider  $X_1$  symmetric stable of index  $\alpha$ , that is  $\mathbb{E} e^{itX_1} = e^{-|t|^\alpha}$ .

As is well-known  $\mathbb{E}|X_1|^p < \infty \forall p < \alpha$ ,  $\mathbb{E}|X_1|^\alpha = \infty$  and the weak law holds iff  $\alpha > 1$  (since it is equivalent to the Fourier transform diff. at 0, and the deriv. is the value that  $\frac{S_n}{n}$  is close to - theorem of E. Pitman).

for  $\alpha = 1$ , this is the Cauchy dist., with density  $\frac{1}{\pi(1+x^2)}$ .

From the Fourier trans. criterion we will mention, these walks are neighbourhood-recurrent iff  $\alpha \geq 1$  (with  $\alpha = 1$  included)

5.11 (Thepp 64):  $\exists$  RW with arbitrarily large tails which are point-recurrent, that is,  $\forall \epsilon(x) \rightarrow 0$  as  $x \rightarrow \infty$   $\exists X_1$  s.t.

$P(X_1 > x) \geq \epsilon(x)$  for large  $x$ , and  $S_n$  point-recurrent.

Give as exercise with hint.

correction from last class: we showed for SRW, if  $k = o(n^{3/4})$

has the same parity as  $n$  then  $P(S_n = k) \sim \sqrt{\frac{2}{\pi n}} e^{-k^2/2n}$  as  $n \rightarrow \infty$ .

In fact, we had this estimate uniformly in  $k$  for  $k = o(n^{3/4})$ .

Lemma: For  $k = o(n^{3/4})$ ,  $P(S_n \geq k) \geq c \frac{\sqrt{n}}{k} e^{-k^2/2n}$  if  $k > \sqrt{n}$ .

Proof:  $P(S_n \geq k) \geq P(k \leq S_n \leq k + (\frac{n}{k} \wedge \sqrt{n})) \geq \frac{c}{\sqrt{n}} \sum e^{-j^2/2n}$

noting that for these  $j$ , if  $k \leq \sqrt{n}$ , we have  $e^{-j^2/2n} \geq e^{-(k+\sqrt{n})/2n} = e^{-\frac{k^2}{2n} - \frac{k}{\sqrt{n}} - \frac{1}{2}} \geq c e^{-k^2/2n}$

similarly, if  $k > \sqrt{n}$ ,  $e^{-j^2/2n} \geq e^{-(k+\frac{n}{k})/2n} = e^{-\frac{k^2}{2n} - 1 - \frac{n}{2k^2}} \geq c e^{-\frac{k^2}{2n}}$ .

Thus  $P(S_n \geq k) \geq \frac{c}{\sqrt{n}} e^{-k^2/2n} \cdot (\frac{n}{k} \wedge \sqrt{n}) \geq \begin{cases} c & \text{if } k < \sqrt{n} \text{ (also follows from SLLN)} \\ c \frac{\sqrt{n}}{k} e^{-k^2/2n} & \text{if } k > \sqrt{n} \end{cases}$

In proof of LIL, we had  $A_k = \{S_{a^k} - S_{a^{k-1}} \geq (1-\delta)U(a^k - a^{k-1})\}$

for  $U(t) = \sqrt{2t \log \log t}$ ,  $0 < \delta < 1$  and  $a > 1$ , and we needed

to show  $\sum_{k=1}^{\infty} P(A_k) = \infty$ . Indeed, by prev. lemma for large  $k$ :

$$P(A_k) \geq c \frac{\sqrt{a^k - a^{k-1}}}{(1-\delta)U(a^k - a^{k-1})} e^{-\frac{(1-\delta)^2 U^2(a^k - a^{k-1})}{2(a^k - a^{k-1})}} =$$

$$= C' \frac{1}{(1-\delta)\sqrt{\log \log(a^k - a^{k-1})}} \frac{1}{[\log(a^k - a^{k-1})]^{(1-\delta)^2}}$$

which is not summable in  $k$ .

Proof of recurrence thm. in  $d=2$ :

Let  $G(x, y) = E^x \# \text{visits to } y = \sum_{n=0}^{\infty} P^x(S'_n = y)$ .

Lemma:  $G(x, y) = P^x(\text{Visit } y) \cdot G(x, x) \leq G(x, x)$ .

strong Markov Property

Proof: Let  $T = \min\{n / S'_n = y\}$ . Then

$$G(x, y) = \sum_{n=0}^{\infty} P^x(S'_n = y) = \sum_{n=0}^{\infty} \sum_{k=0}^{\infty} P^x(S'_n = y, T = k) = \sum_{k=0}^{\infty} \sum_{n=0}^{\infty} P^x(S'_n = y, T = k) =$$

$$= \sum_{k=0}^{\infty} \sum_{n=0}^{\infty} P^x(T = k) \cdot P^y(S'_n = y) = \sum_{k=0}^{\infty} P^x(T = k) \underbrace{G(y, y)}_{= G(x, x)} = P^x(\text{Visit } y) G(x, x)$$

Proof of thm.: need to show that  $\sum_{n=0}^{\infty} P(S'_n = 0) = G(0, 0) = \infty$ .

Have  $\forall b > 0$ ,  $P(|S'_n| \leq b) \rightarrow \int_{|x| \leq b} n(x) dx$  where  $n$  is the normal density.

Note that if we had an LCLT of form  $P(S'_n = 0) \geq \frac{c}{n}$ , we would be done.

The idea is to get something similar from the CLT.

By the prev. lemma:  $\sum_{n=0}^{\infty} P(S'_n = 0) \geq \frac{c}{m^2} \sum_{n=0}^{\infty} P(|S'_n| \leq m)$

if we had,  $P(S'_n = 0) \geq \frac{1}{n} P(|S'_n| \leq m)$ , it would also suffice.

$P(|S'_n| \leq m) = \int_{-m}^m n(x) dx \geq \int_{-m}^m \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \geq \frac{1}{\sqrt{2\pi}} \int_{-m}^m \frac{1}{2} dx = \frac{1}{\sqrt{2\pi}} m$

Note  $\frac{1}{m^2} \sum_{n=0}^{\infty} P(|S'_n| \leq m) = \int_0^{\infty} \int_{-m}^m P(|S'_\theta| \leq m) d\theta$ . By Fatou:  $\liminf \frac{1}{m^2} \sum_{n=0}^{\infty} P(|S'_n| \leq m) \geq \int_0^{\infty} \int_{-m}^m n(x) dx d\theta \geq \int_0^{\infty} \int_{-m}^m \frac{1}{2} dx d\theta = \infty$ . Thus  $\sum_{n=0}^{\infty} P(S'_n = 0) = \infty$ .

Fourier analytic criterion for recurrence: (for RW on  $\mathbb{Z}^d$ )

Thm: Let  $\varphi(\theta) = E e^{i\theta \cdot X_1}$  be the char. fcn. of the RW.

Then  $S_n$  is recurrent iff  $\lim_{r \uparrow 1} \int_{[-\pi, \pi]^d} \frac{1}{1 - r\varphi(\theta)} d\theta = \infty$

Remark: A much more difficult thm. is that the prev.

cond. can be replaced by  $\int_{[-\pi, \pi]^d} \frac{1}{1 - \varphi(\theta)} d\theta = \infty$ .

2) A similar statement is true for RW on  $\mathbb{R}^d$ . There

the walk is neighbourhood-recurrent iff for some  $d > 0$  (then for all  $d$ ) we have  $\int_{(-d, d)^d} \frac{1}{1-\varphi(\theta)} d\theta = \infty$  (or the statement with  $r$ )

Proof: recall that  $\varphi^n(\theta) = \mathbb{E} e^{i\theta \cdot S_n}$  and by the Inv. formula  $P(S_n = y) = (2\pi)^{-d} \int_{[-\pi, \pi]^d} e^{-iy \cdot \theta} \varphi^n(\theta) d\theta$ .

Thus  $P(S_n = 0) = (2\pi)^{-d} \int_{[-\pi, \pi]^d} \varphi^n(\theta) d\theta$  we have

$$\sum_{n=0}^{\infty} r^n P(S_n = 0) = \sum_{n=0}^{\infty} (2\pi)^{-d} \int_{[-\pi, \pi]^d} r^n \varphi^n(\theta) d\theta = (2\pi)^{-d} \int_{[-\pi, \pi]^d} \frac{1}{1-r\varphi(\theta)} d\theta$$

abs. summable

Since the LHS is real

$$\sum_{n=0}^{\infty} r^n P(S_n = 0) = (2\pi)^{-d} \int_{[-\pi, \pi]^d} \operatorname{Re} \left[ \frac{1}{1-r\varphi(\theta)} \right] d\theta$$

and we conclude since  $\lim_{r \uparrow 1} \sum_{n=0}^{\infty} r^n P(S_n = 0) = \sum_{n=0}^{\infty} P(S_n = 0)$

### Lecture 5 28.3.11

(scribe) Give proof from prev. page of 2D Chung-Fuchs thm.

Remind Fourier analytic recurrence criterion.

Perhaps mention that in 2D, criterion applied to sym.

stable dist. with char. fcn.  $\varphi(t) = e^{-|t|^\alpha}$   $t \in \mathbb{R}^2, |t| = \sqrt{t_1^2 + t_2^2}$  ( $\alpha \leq 2$ )

is neighbourhood-recurrent only in the Gaussian case  $d=2$ .

Def.: A RW in 3D is truly 3-D if  $P(X_1 \neq 0) > 0 \forall t \neq 0$ .  
(Does not stay in a plane)

Proof of 3D Chung-Fuchs thm.: Idea: in 1D  $\mathbb{E} e^{itX_1} = 1 + o(t^2)$  ( $t \rightarrow 0$ )  
(based on Durrett Chap. 4) then  $X_1 \equiv 0$  since it implies first and second moments are 0.

In more than 1D, same applies in each direction  $\varphi(t) = \mathbb{E} e^{i(t \cdot X_1)}$

1)  $\operatorname{Re} \frac{1}{1-z} \leq \frac{1}{\operatorname{Re}(1-z)}$  if  $\operatorname{Re}(z) \leq 1$ . Proof: Write  $1-z = re^{i\theta}$  with  $\theta \in [-\pi/2, \pi/2]$ .

Then  $\frac{1}{1-z} = \frac{1}{r} e^{-i\theta}$ ,  $\operatorname{Re} \left( \frac{1}{1-z} \right) = \frac{1}{r} \cos \theta$ ,  $\operatorname{Re}(1-z) = r \cos \theta$ ,  $\frac{1}{\operatorname{Re}(1-z)} = \frac{1}{r \cos \theta}$  Assume  $|\theta| < \pi/2$  by contin.

Need  $\cos \theta \leq \frac{1}{\cos \theta}$ . True if  $\cos \theta > 0$ .

2)  $1 - \cos(x) \geq x^2/4$  for  $|x| \leq \pi/2$ . Proof:  $1 - \cos x = \int_0^x \sin y dy = \int_0^x \int_0^y \cos z dz \geq \int_0^x \int_0^y \frac{1}{2} dz dy = \frac{x^2}{4}$ .

3)  $\operatorname{Re}(1 - \mathbb{E} e^{it \cdot X_1}) = \mathbb{E}(1 - \cos(t \cdot X_1)) \geq \mathbb{E} \left[ \frac{1}{4} \mathbb{1}_{|t \cdot X_1| \leq \pi/2} \right]$

Write  $t = ru$  with  $r > 0$  and observe  

$$P\{1 - \varphi(ru) \geq \frac{r^2}{4} \mathbb{E} \left( \mathbb{1}_{\left| \frac{r}{|u \cdot X_1|} \leq \frac{\pi}{3} \right.} \cdot (u \cdot X_1)^2 \right)$$

4) By continuity of  $\mathbb{E}(u \cdot X_1)^2$  in  $u$  and compactness of  $S^2$  (not needed if  $\mathbb{E}(u \cdot X_1)^2 = \min_{u \in S^2} \mathbb{E}(u \cdot X_1)^2 > 0$  since  $X_1$  is truly 3-D.)

Furthermore, since if  $r \rightarrow 0$  and  $u(r) \rightarrow u$  we have by Fatou

$$\liminf_{r \rightarrow 0} \mathbb{E} \left( \mathbb{1}_{\left| \frac{r}{|u \cdot X_1|} \leq \frac{\pi}{3} \right.} (u \cdot X_1)^2 \right) \geq \mathbb{E}(u \cdot X_1)^2 > 0 \text{ since } X_1 \text{ is truly 3-D.}$$

Hence by compactness  $\exists r_0$  s.t.  $\forall r < r_0, u \in S^2, \mathbb{E} \left( \mathbb{1}_{\left| \frac{r}{|u \cdot X_1|} \leq \frac{\pi}{3} \right.} (u \cdot X_1)^2 \right) > c > 0$

5) Thus (3) and (4) give  $P\{1 - \varphi(ru)\} \geq cr^2$  for  $r < r_0$ .

6) It follows that:

$$\int_{\pi, \pi]^3} \frac{1}{|t|} dt \geq \int_{\pi, \pi]^3} \frac{1}{|t|} dt \geq \int_{|t| < r_0} \frac{1}{|t|^2} dt = \int_0^{r_0} \frac{r^2}{cr^2} dr < \infty, \text{ implying the walk is transient.}$$

(Note that in d-dim., result follows from 3D result, but also get  $\int_0^{r_0} \frac{r^{d-1}}{cr^2} dr$  in the calculation for a truly d-dim. walk.)

martingales (need conv.) and coupling trivial on reg. conn. graphs. harmonic fns. and tail  $\sigma$ -field. Demante (in discrete 6.2) (Lewy 1.4) harmonic fns. off a set in  $\mathbb{Z}^d$ .

Rainanovich-Vershik - Sublinear harmonic fns., non-NEG. harmonic fns. Liouville Poisson boundary

review of martingales in discrete time

DEF.: Filtration  $\mathcal{F}_0 \subset \mathcal{F}_1 \subset \dots$   $\sigma$ -fields. equality mod 0.   
 Martingale:  $\{M_n\}_{n \geq 0}$  seq. of integrable RV's with  $M_n$  wrt. filtration  $\{\mathcal{F}_n\}_{n \geq 0}$  meas. wrt.  $\mathcal{F}_n$  and  $\forall m \leq n, M_m = \mathbb{E}(M_n | \mathcal{F}_m)$    
 supermart.:  $\forall m \leq n, M_m \geq \mathbb{E}(M_n | \mathcal{F}_m)$ , submart.:  $\forall m \leq n, M_m \leq \mathbb{E}(M_n | \mathcal{F}_m)$

sufficient to check:  $\forall n, M_n = \mathbb{E}(M_{n+1} | \mathcal{F}_n)$    
 If the filtration is not specified then  $\mathcal{F}_n = \sigma(M_0, \dots, M_n)$ .

If  $M_0, X_1, X_2, \dots$  ind.,  $\mathbb{E}|M_0| < \infty$  and  $\mathbb{E}X_i = 0 \forall i$  then  $M_n = M_0 + X_1 + \dots + X_n$  is a mart.

Prop: Jensen's ineq. gives that if  $f: \mathbb{R} \rightarrow \mathbb{R}$  is convex then if  $\{M_n\}_{n \geq 0}$  is a mart. and  $\mathbb{E}|f(M_n)| < \infty \forall n$  then  $\{f(M_n)\}$  is a submart.

In particular, can take  $\{M_n\}^\alpha$  for  $\alpha \geq 1$  and  $e^{bM_n}$  for  $b \in \mathbb{R}$  and get a submart. (assuming these are int.)

Optional sampling

Def.: Stopping time  $T$  wrt. the filtration  $\{\mathcal{F}_n\}_{n \geq 0}$  is a  $\{0, 1, \dots\} \cup \{\infty\}$  valued RV s.t.  $\forall n, \{T \leq n\}$  is  $\mathcal{F}_n$ -meas.

Note that the min. of two stop. times is a stop. time and that const. are stop. time.

Prop.:  $\{M_n\}_{n \geq 0}$  mart.,  $T$  stop. time  $\Rightarrow \{M_{T \wedge n}\}_{n \geq 0}$  is a mart.

In part.,  $E\{M_T\} = E\{M_0\}$ . Same for submartingale or superm.

Proof: Immediate from def.  $M_{T \wedge (n+1)} = M_{T \wedge n} \mathbb{1}_{\{T \leq n\}} + M_{n+1} \mathbb{1}_{\{T > n+1\}}$

Thm. (optional sampling thm.):  $\{M_n\}_{n \geq 0}$  mart. and  $T$  stop. time wrt.  $\{\mathcal{F}_n\}_{n \geq 0}$

Assume  $P(T < \infty) = 1$  and at least one of the following holds:

1)  $T$  is bdd. (i.e.,  $P(T \leq K) = 1$  for some  $K$ ).

2) (Dom. conv.)  $\exists Y \in L_1$  s.t.  $|M_T| \leq Y$ .

3)  $E|M_T| < \infty$  and  $\lim_{n \rightarrow \infty} E(|M_n| \mathbb{1}_{\{T > n\}}) = 0$ .

4)  $\{M_n\}$  are unif. int. i.e.,  $\forall \epsilon > 0 \exists K_\epsilon$  s.t.  $E(|M_n| \mathbb{1}_{\{M_n > K_\epsilon\}}) < \epsilon \forall n$ .

5)  $\exists \alpha > 1$  s.t.  $\{M_n\}_{n \geq 0}$  are bdd. in  $L_\alpha$  ( $E|M_n|^\alpha \leq K \forall n$ ).

Then  $E M_0 = E M_T$  (in part.,  $M_T \in L_1$ ). For submart.,  $E M_0 \leq E M_T$ .

Proof: 1, 2 clear. 5  $\Rightarrow$  4. For 3,  $M_T = M_{T \wedge n} + M_T \mathbb{1}_{\{T > n\}} - M_n \mathbb{1}_{\{T > n\}}$ .

Since  $\mathbb{1}_{\{T > n\}} \rightarrow 0$  d.s. and  $E|M_T| < \infty$ , dom. conv. gives

$\lim_{n \rightarrow \infty} E M_T \mathbb{1}_{\{T > n\}} = 0$ . Hence 3 is suff.  $\uparrow$  (A14 in Williams),  $\uparrow \Rightarrow$  3? (Lower limit)

6:  $|M_{T \wedge n} - M_0| = |\sum_{k=1}^{T \wedge n} (M_k - M_{k-1})| \leq K T$ . Dom. conv. gives  $E(M_T - M_0) = 0$

Doob's maximal inequality: If  $\{M_n\}_{n \geq 0}$  is a nonneg. submart. wrt.  $\{\mathcal{F}_n\}$  and  $\lambda > 0$

then  $P(\max_{0 \leq j \leq n} M_j \geq \lambda) \leq \frac{E M_n}{\lambda}$ .

Especially useful as: If  $\{M_n\}$  is a mart.,  $\alpha \geq 1, b \geq 0$ :

$$P(\max_{0 \leq j \leq n} M_j \geq \lambda) \leq \frac{E|M_n|^\alpha}{\lambda^\alpha}$$

$$P(\max_{0 \leq j \leq n} M_j \geq \lambda) \leq \frac{E e^{bM_n}}{e^{b\lambda}}$$

Martingale conv. thm.:  $\{M_n\}_{n \geq 0}$  supermart.,  $\sup_n E(M_n) < \infty$  then (eg.  $M_n \geq 0$ )

$\exists$  RV  $M$  s.t.  $M_n \rightarrow M$  a.s. Convergence need not occur in  $L_1$ . However,  $M \in L_1$  (Durrett 5.2.8) eg. RW stopped

2) For a submart., TFAR: 1)  $\uparrow$  or 2) conv. a.s. and in  $L_1$ , 3) conv. in  $L_1$ . When hitting 1.

For a mart. also: 1)  $\uparrow$   $\{M_n\} \in L_1, M_n = E(M | \mathcal{F}_n)$

Williams:  $\triangleright$  signifies something important  $\triangleright\triangleright$  something very important  $\triangleright\triangleright\triangleright$  the mart. conv. thm.

Cor.: If  $\mathcal{F}_n \uparrow \mathcal{F}_\infty$  (in part.,  $\mathcal{F}_\infty = \sigma(\cup \mathcal{F}_n)$ ) then  $\forall X \in L$ ,

$$\mathbb{E}(X | \mathcal{F}_n) \xrightarrow{n \rightarrow \infty} \mathbb{E}(X | \mathcal{F}_\infty) \text{ a.s. and in } L_1.$$

Case  $X = 1_A$  called Lévy's 0-1 law. Generalizes Kolmogorov 0-1 law.

### Harmonic Functions

A function  $h: G \rightarrow \mathbb{R}$  (loc. fin. (always assumed)) is called harmonic if  $\forall x$

$$h(x) = \frac{1}{|N(x)|} \sum_{y \in N(x)} h(y) \text{ where } N(x) = \{y | y \sim x\}. \text{ In other words,}$$

if  $\{Z_n\}_{n \geq 0}$  is a SRW on  $G$ , for any starting vertex  $x$ ,  $h(x) = \mathbb{E}^x h(Z_1)$

Thus if  $h(Z_n)$  is int.  $\forall n$ , then  $\{h(Z_n)\}_{n \geq 0}$  is a mart.

DEF.: The space of all bdd harmonic fcn's on  $G$  is called the Poisson boundary of  $G$ .  $G$  is called Liouville if the only bdd. harmonic fcn's on it are constants.

Remark. Usually studied for the Cayley graphs of groups. Sometimes harmonicity is defined wrt. other RW's.

Prop.: On any conn. recurrent graph, any non-neg. harmonic fcn. is constant.

Proof: Fix  $x, y \in G$  and a non-neg. harmonic fcn.  $h$ . Let  $M_n := h(Z_n)$

With  $Z_0 = x$ .  $M_n$  is bdd. and hence int. since  $G$  is loc. fin. By the mart. conv. thm,  $M_n \rightarrow M$  a.s.

Since  $G$  is recurrent,  $\mathbb{P}(Z_n = x \text{ i.o.}) = \mathbb{P}(Z_n = y \text{ i.o.}) = 1$ .

Hence  $\mathbb{P}(M = h(x)) = \mathbb{P}(M = h(y)) = 1$ . Thus  $h(x) = h(y)$ .

DEF.: A coupling of two dist.  $\mu, \nu$  is a joint dist.  $(X, Y)$

With  $X \sim \mu$  and  $Y \sim \nu$ .

Aldous-Thorisson (93) paper, Lindvall (92), Thorisson book, Karlin-Vershik (83)

Lecture 6, 4.4.11

Scribe HW

Prop.: If for any  $x, y \in G \exists$  coupling of the SRW

$Z_n^x, Z_n^y$  with  $Z_0^x = x, Z_0^y = y$  s.t.  $\mathbb{P}(\exists K \text{ for which } Z_n^x = Z_{n+K}^y \text{ for inf. many } n) = 1$

then any bdd. harmonic fcn. on  $G$  is constant.

Proof:  $M_n^x := h(Z_n^x), M_n^y := h(Z_n^y)$  are mart. conv. to  $M^x, M^y$  resp.

By assumption, under our coupling  $\mathbb{P}(M_n^x = M_{n+K}^y \text{ for inf. many } n) = 1$  and some  $K$

Hence  $\mathbb{P}(M^x = M^y) = 1$ . Since  $M_n^x$  and  $M_n^y$  are bdd. and hence  $\mathbb{E} M_n^x = \mathbb{E} M_n^y$

We get  $h(x) = \mathbb{E} M^x = \mathbb{E} M^y = h(y)$ .

Examples on  $\mathbb{Z}^d$ : constants, linear fcn's.

Got to here but also showed example of bdd. harmonic fcn. on tree

Shift-coupling with Peter in context of Markov chains. Same as two times (T, T) s.t.  $Z_t^x = Z_t^y$  and  $\mathbb{P}(T < \infty) = 1$

Prop.: on  $\mathbb{Z}^d$  there are no odd harmonic fcn's.

Proof: It is sufficient to find a shifted coupling for any two starting points  $x, y \in \mathbb{Z}^d$ . First assume  $x-y$  has all even coord. Do the following coupling, at each step  $z_n^x$  and  $z_n^y$  move in the same coord., ind. if that coord. is diff. in  $z_n^x$  and  $z_n^y$  and together others. Since 1D SRW is recurrent, this gives a shifted coupling (with no shift). If not all coords. of  $x-y$  are even, run  $z_n^x$  until time  $K$  when  $z_n^x - y$  has all coords. even, then do prev. coupling. Since  $P(K < \infty) = 1$ , this gives a shifted coupling with shift  $K$ .

### Harmonic fcn's, $\sigma$ -fields and coupling

We describe the study of harmonic fcn's. in the wider context of Markov chains, though our main example will still be RW on graphs.

Let  $(E, \mathcal{E})$  be a countable state space,  $p: S \times S \rightarrow [0, 1]$  a Markov kernel ( $\sum_y p(x, y) = 1 \forall x$ ), the space of trajectories is  $S^{\mathbb{N}}$  with the product  $\sigma$ -algebra. DEF:  $\theta: S^{\mathbb{N}} \rightarrow S^{\mathbb{N}}$ ,  $\theta((x_n)_{n \geq 0}) = (x_{n+1})_{n \geq 0}$ , the shift.

$\forall \tau \in \mathcal{T}$  RV,

a fcn. of any tail of the trajectory

Tail  $\sigma$ -field:  $\mathcal{T} = \bigcap_{n=0}^{\infty} \theta_n^{-1}(\mathcal{E}^{\infty})$ , inside the intersection is the  $\sigma$ -field of events depending only on coords. from  $n$  onward.

$\forall \mathcal{I} \in \mathcal{I}$  RV, a fcn. of the traj. inv. under shifts

Invariant  $\sigma$ -field:  $\mathcal{I} = \{B \in \mathcal{E}^{\infty} \mid \theta_1^{-1} B = B\}$ , all events holding for a trajectory iff they hold for any shift of it.

note:  $\mathcal{I} \subseteq \mathcal{T}$  as can be easily verified.

Examples: 1) SRW on  $\mathbb{Z}^d$ : the parity of the starting point  $X_0$  can be deduced from observing  $X_n$  for any  $n$ . Hence if the initial dist.  $\mu$  is supp. on both even and odd points then this event has non-trivial prob. under  $\mathbb{P}^{\mu}$ .

2) In a periodic MC, the initial period is in the tail  $\sigma$ -field.

3) SRW on a complete binary tree: The walk eventually fixates on some end. The event that it finishes in the bottom half is invariant under shifts and hence belongs to  $\mathcal{I}$ . It has non-trivial prob. under any  $\mathbb{P}^x$ . Here  $\text{coo, parity}$  is in the tail  $\sigma$ -field.

DEF: Identification mod 0: We shall identify two RV's  $U, V \in \mathcal{E}^{\infty}$  if for any  $x \in E$   $\mathbb{P}^x(U \neq V) = 0$ . We may write  $U = V \pmod{0}$ .

The following is the fundamental relation between these  $\sigma$ -fields and bounded harmonic fcn's.

Def: A p-harmonic fcn.  $h: E \rightarrow \mathbb{R}$  is a fcn. satisfying  $h(x) = \sum_{y \in E} h(y) P(x, y) \quad \forall x \in E$ .

A p-spacetime-harmonic fcn.  $h: E \times \mathbb{Z} \rightarrow \mathbb{R}$  is a fcn. satisfying

$$h(x, n) = \sum_{y \in E} h(y, n+1) P(x, y) \quad \forall x \in E, n \geq 0. \quad \leftarrow \text{Example: } \left( \frac{z_n}{n} \right)$$

Thm. 1) There is a bijection between bounded  $\mathcal{U} \in \mathcal{T}$  (Parity of  $z_n$ , say.

considered modulo and bounded spacetime harmonic fcn's.

Given by  $\mathcal{U} \in \mathcal{T} \Rightarrow h(x, n) = \mathbb{E}^x(\mathcal{U}_n)$  where  $\mathcal{U} = \mathcal{U}_n \circ \theta^n$ .

2) There is a bijection between bounded  $\mathcal{U} \in \mathcal{X}$  modulo and

bdd. harmonic fcn's,  $\mathcal{U} \in \mathcal{X} \Rightarrow h(x) = \mathbb{E}^x \mathcal{U}$ .

Proof 1)  $\Rightarrow$ :  $\mathcal{U} \in \mathcal{T}$ . Define  $h(x, n) = \mathbb{E}^x(\mathcal{U}_n)$  where  $\mathcal{U} = \mathcal{U}_n \circ \theta^n$ .

$h$  is clearly bdd.  $h(x, n) = \mathbb{E}^x(\mathcal{U}_n) = \mathbb{E}^x \mathbb{E}^x(\mathcal{U}_n | \mathcal{Z}_n) = \mathbb{E}^x \mathbb{E}^{\mathcal{Z}_n}(\mathcal{U}_{n+1}) = \mathbb{E}^x h(\mathcal{Z}_n, n+1)$

So that  $h$  is spacetime harmonic.

$\Leftarrow$ :  $h(x, n)$  bdd. spacetime harmonic. For each  $x$ ,  $(h(\mathcal{Z}_n^x, n))_{n \geq 0}$  is a bdd. martingale and hence converges. The limit is  $\mathcal{U}$  under  $\mathbb{P}^x$ .

By def. as a limit, it is a tail RV.

Finally, since  $h$  is bdd., we also have  $h(\mathcal{Z}_n^x, n) = \mathbb{E}^x(\mathcal{U} | \mathcal{F}_n)$ .

So that  $\mathcal{U}$  is indeed the limit of  $h$  and the map  $\mathcal{U} \rightarrow h$  is 1-1.

2) Similar proof.  $h(x) = \mathbb{E}^x \mathcal{U} = \mathbb{E}^x \mathbb{E}^x(\mathcal{U} | \mathcal{Z}_n) = \mathbb{E}^x \mathbb{E}^{\mathcal{Z}_n} \mathcal{U} = \mathbb{E}^x h(\mathcal{Z}_n)$ .

Couplings: Prop. 1) If  $\forall x, y \in E \exists$  coupling of  $\mathcal{Z}_n^x, \mathcal{Z}_n^y$  with  $T := \min\{n | \mathcal{Z}_n^x = \mathcal{Z}_n^y\}$  satisfying  $\mathbb{P}(T < \infty) = 1$  (a successful coupling) then all bdd.

spacetime harmonic fcn's are constant.

Proof: Fix  $x, y$ .  $h(\mathcal{Z}_n^x, n)$  and  $h(\mathcal{Z}_n^y, n)$  are bdd. martingales.

$$|h(x, 0) - h(y, 0)| = |\mathbb{E} h(\mathcal{Z}_n^x, n) - \mathbb{E} h(\mathcal{Z}_n^y, n)| = |\mathbb{E} (h(\mathcal{Z}_n^x, n) - h(\mathcal{Z}_n^y, n)) \mathbb{1}_{\{T > n\}}| \leq 2M \mathbb{P}(T > n) \text{ where } M = \max_{z, n} |h(z, n)|. \text{ Hence } h(x, 0) = h(y, 0) \quad \forall x, y.$$

Similarly using that  $h(\mathcal{Z}_n^x, n+m)$  is a mart. for every  $n$  gives  $h(x, m) = h(y, m) \quad \forall x, y, m$ .

Prop. 2) If  $\forall x, y \in E \exists$  coupling of  $\mathcal{Z}_n^x, \mathcal{Z}_n^y$  s.t.  $\mathbb{P}(\exists k \text{ s.t. } \mathcal{Z}_{n+k}^x = \mathcal{Z}_{n+k}^y \text{ for all large } n) = 1$  (a successful shift-coupling) then all bdd. harmonic fcn's are constant.

Then give thm. with many equiv. and then comp.

Then I aim: Vershik (ask Gili)

Finally Lawler for Dirichlet Problem.

proof: Similar to prev. case.

In fact, these statements are iff and we have many other equivalences:

Thm.: The following are equivalent for a Markov chain

- |   |   |
|---|---|
| 1) Trivial tail $\sigma$ -field for every initial distribution.   | 1) Trivial invariant $\sigma$ -field for any initial dist.  |
| 2) All bdd. spacetime harmonic fcts. are constant.  | 2) All bdd. harmonic fcts. are constant.  |
| 3) $\exists$ a successful coupling of $Z_n^x, Z_n^y \forall x, y$ .   | 3) $\exists$ a successful shift-coupling of $Z_n^x, Z_n^y \forall x, y$ .   |
| 4) The Markov chain is mixing.  | 4) The Markov chain is Cesaro-mixing.   |
| 5) $\ P^n - P'^n\ _{TV} \xrightarrow{n \rightarrow \infty} 0 \forall$ initial $\mu, \mu'$   | 5) $\frac{1}{t} \left\  \sum_{s=0}^t P^s - \sum_{s=0}^t P'^s \right\ _{TV} \xrightarrow{t \rightarrow \infty} 0 \forall \mu, \mu'$              |
| 6) $\ P^{(0^n)} Z(\cdot) - P^{(0'^n)} Z(\cdot)\ _{TV} \xrightarrow{n \rightarrow \infty} 0 \forall \mu, \mu'$<br>measures on processes. | 6) $\ P^{(0^n)} Z(\cdot) - P^{(0'^n)} Z(\cdot)\ _{TV} \xrightarrow{n \rightarrow \infty} 0$<br>where $0 \sim \text{unif}(0,1)$ is ind. of $Z$ . |

proof of (5) From (3) (3) gives a coupling for any initial dist.

Allows to prove rate of conv. results!

$$|P^n(Z_n \in A) - P'^n(Z_n \in A)| = \left| E \left[ \mathbb{1}_{(Z_n \in A)} - \mathbb{1}_{(Z_n' \in A)} \mid T > n \right] \right| \leq P(T > n).$$

$$d_{TV}(v, v') = \max_A |v(A) - v'(A)|$$

(The coupling inequality)

For the other direction, it is possible to show that there exists a coupling for which equality is attained above for all  $n$  (maximal coupling)

EXAMPLES: 1)  $Z^0$ : Tail  $\sigma$ -field is only parity and invariant field is trivial two

Later, extensions to sub-invar and reg-ns

Proof: coupling from prev. page shows that if  $x$  and  $y$  have same parity then  $h(x, m) = h(y, m)$  since there exists a

Exercise: prove tail is just parity using a coupling

successful coupling. Hence  $h$  is just parity.

Similarly,  $\forall x, y \exists$  successful shift-coupling. (no. of possible return times is  $J$ )

2) An irreducible recurrent chain with all states of period  $J$ .

$J = \{ \text{cyclic class of starting point} \}$ ,  $\neq$  trivial / we proved this last section (or see Durrett for a proof). (last time can get from any state to any state) (period  $J$ )

3) Long range 1D RW, Ornstein's coupling: If  $\nu_n$  is a 1D irreducible, aperiodic

on  $\mathbb{Z}$  RW then  $\tau$  is trivial. Fix  $x, y \in \mathbb{Z}$  and let  $(X_n)_{n \geq 0}$  be a seq. of IID incr.

Take  $M$  large enough so that the walk with incr.  $(X_n | X_n \leq M)$  is

still irreducible and aperiodic. Let  $(X_n^M)_{n \geq 0}$  be ind. of  $(X_n)$  and IID

dist. as  $(X_n | X_n \leq M)$ .

Set  $S_n^x = S_{n-1}^x + X_{n-1}$

Now couple  $S_n^x$  and  $S_n^y$  by: IF  $S_{n-1}^x = S_{n-1}^y$  then  $S_n^x = S_n^y$   
 otherwise, if  $X_n > M$  then  $S_n^y = S_{n-1}^y + X_{n-1}$ . IF  $X_n \leq M$   
 then  $S_n^y = S_{n-1}^y + X_{n-1}^M$  (big jumps together and small jumps ind.)  
 since  $(S_n^y - S_n^x)_{n \geq 0}$  is a 1D RW on  $\mathbb{Z}$ , started at  $x-y$ , with  
 mean 0, and irreducible <sup>check!</sup> it follows that it will hit 0 eventually.

Hence this is a successful coupling.

needed?

4) RW on a tree:  $\mathcal{X}$  is the  $\sigma$ -field generated by the events  
 of going to infinity on a given end.  $\tau$  adds to that  
 the parity of the starting point  $\epsilon_0$

Kaim. Vershik  
 Lawler

(1983) Kaimanovich-Vershik thm RW  $(S_n)_{n \geq 0}$  on the Cayley graph of  
 a (discrete, countable) group with finite entropy increments  
 has no non-constant h.d harmonic fcn's. iff.  $h(G, \nu) := \lim_{n \rightarrow \infty} \frac{H(S_n)}{n} = 0$

where  $H(x) := -\sum p^i(x) \log(p^i(x))$  (with  $0 \log 0 := 0$ )

lamblihter jumps

5) Lamblihter groups:  $\mathbb{Z}_2 \wr G$ : Cayley graph indexed by  $(g, f: G \rightarrow \mathbb{Z}_2)$ .

Liouville iff  $G$  is recurrent.

IF  $G$  is transient, <sup>limit</sup> state of lamp at e.g. is an invariant event.

IF  $G$  is recurrent,  $\frac{R_n}{n} \rightarrow 0$  by KSW thm,  
 $R_n$  ← range of SRW on  $G$

check!

From this one can deduce that  $H(S_n) = o(n)$  on  $\mathbb{Z}_2 \wr G$ .

so finish by Kaimanovich-Vershik.

Additional characterization: srw on Cayley graph Liouville  $\Leftrightarrow \mathbb{E} \downarrow(S_n, 0) = o(n)$

Open questions 1) IS Liouville a group prop. (preserved  
 under change of gen. in Cayley graph)?

2) Possible rates of escape for SRW on groups?  
 are they preserved under change of gen.?

Lecture 7 11.4.11 Scribe

Harmonic functions on sets in  $\mathbb{Z}^d$ . CLT in  $\mathbb{Z}^d$ .  
 Green's fcn's in  $\mathbb{Z}^d$  and their asymptotics. Number of values seen by  
 a RW before a geom. time variant of  $K_{\text{RW}}$

Martin capacity for Markov chains.  
 Intersection of RW's  
 Support of harmonic measure in  $\mathbb{Z}^d$  (paper of Itai)

# Introduction to Potential theory (Lawler-Limic, Lawler)

## 1. Harmonic fcn. on $\mathbb{Z}^d$ and the Dirichlet Problem

$$\frac{1}{2} \sum_{y \sim x} f(y) = \frac{1}{2} \sum_{y \sim x} [f(y) - f(x)] + f(x) = \mathbb{E}^x f(S_1) - f(x)$$

Accumulate Notation on side-board

For  $f: \mathbb{Z}^d \rightarrow \mathbb{R}$ , define  $(\mathcal{L}f)(x) = \frac{1}{2} \sum_{y \sim x} [f(y) - f(x)] = \frac{1}{2} \sum_{y \sim x} f(y) - f(x) = \mathbb{E}^x f(S_1) - f(x)$

the discrete Laplacian,  $f$  is harmonic iff  $(\mathcal{L}f)(x) = 0$ , subharmonic iff  $(\mathcal{L}f)(x) \geq 0$  (generator of the walk)

For  $A \subseteq \mathbb{Z}^d$ , let  $\tau_A = \min\{n \geq 1 : S_n \notin A\}$ ,  $\bar{\tau}_A = \min\{n \geq 0 : S_n \notin A\}$ .

$\partial A = \{y \in \mathbb{Z}^d - A \mid y \sim x \text{ for some } x \in A\}$  the outer bdr. of  $A$ ,  $\bar{A} := A \cup \partial A$ .

note if  $f: \bar{A} \rightarrow \mathbb{R}$  then  $(\mathcal{L}f)(x)$  is well-defined on  $A$ .

Thm: Suppose  $A \subseteq \mathbb{Z}^d$  satisfies  $P^x(\tau_A < \infty) = 1 \forall x \in A$  and let  $f: \partial A \rightarrow \mathbb{R}$

Other RP and operators are also pos

be a bdd. fcn. Then  $\exists!$  bdd. fcn.  $F: \bar{A} \rightarrow \mathbb{R}$  s.t.

$$(\mathcal{L}F)(x) = 0 \forall x \in A \text{ and } F(x) = f(x) \forall x \in \partial A.$$

It is given by  $F(x) = \mathbb{E}^x F(S_{\bar{\tau}_A})$ . (\*)

Proof: It is simple that (\*) satisfies the requirem. by the Markov Prop.

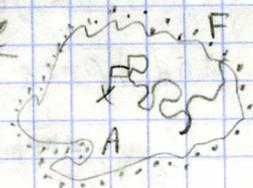
Now if  $F$  is a bdd. fcn. satisfying the requirements

then  $M_n := F(S_{n \wedge \bar{\tau}_A})$  is a bdd. martingale for any starting point.

By the optional sampling thm.  $F(x) = \mathbb{E}^x M_{\bar{\tau}_A} = \mathbb{E}^x F(S_{\bar{\tau}_A})$ .

Remarks: 1) Note the solution satisfies the maximum principle

$$\sup_{x \in \bar{A}} F(x) = \sup_{x \in \partial A} F(x)$$



2) If  $A$  is finite, existence and uniqueness follow by noting there are  $|A|$  equations in  $|A|$  unknowns and the max. principle holds by def. of  $\mathcal{L}$ .

3) If  $|A| = \infty$  there can be inf. many unbdd. sol. e.g.  $A = \{1, 2, \dots\}$ ,  $F(0) = 0$  and  $F(x) = bx$  for some  $b \in \mathbb{R}$ .

4) If  $d = 1, 2$ , the thm. holds  $\forall A$  by recurrence.

On the other hand, if  $d \geq 3$  and  $\exists x \in A$  s.t.  $P^x(\bar{\tau}_A = \infty) > 0$  then  $F(x) = P^x(\bar{\tau}_A = \infty)$  is another bdd. sol. for the boundary data  $f = 0$ .

5) Continuous analogue: If  $D = \{x \in \mathbb{R}^d \mid |x| < 1\}$  and  $f: \partial D \rightarrow \mathbb{R}$  is cont. on  $\bar{D}$ ,

$\Delta^2$  in  $D$  and  $\Delta F(x) = 0 \forall x \in D$  then  $F(x) = \mathbb{E}^x f(B_T)$  for a BM  $B_t$  started at  $x$  and stopped at time  $T = \min\{t \mid |B_t| = 1\}$ .

For  $x \in D$ , the dist. of  $B_T$  (harmonic measure) has a density wrt. Leb. meas. on  $\partial D$  and this density is the Poisson kernel:  $h(x, z) = c_d (1 - |x|^2) / |x - z|^d$ .

Thus  $F(x) = c_d \int_{\partial D} f(z) \frac{1 - |x|^2}{|x - z|^d} dS(z)$  as can be verified directly.

Two facts follow from this repr.: deriv. estimate:  $\forall k \in \mathbb{N}$   $\exists C_k < \infty$  s.t.  $|\nabla^k F(x)| \leq C_k \|f\|_\infty$  for a  $k$ 'th order deriv.  $D^k$ .

Harnack ineq.:  $\forall r \in (0, 1)$  s.t. if  $f \geq 0$  then  $F(x) \leq c_r F(y) \forall x, y$  with  $|x|, |y| \leq r$ .



In other words  $\max_{x \in A} F(x) \leq C$ ,  $\min_{x \in A} F(x)$  for a  $C$ , ind. of  $F$ !

Discrete analogs of these prop. hold but we will not prove them.

What if  $P^x(\tau_A = \infty) > 0$  for some  $x \in A$ ? We will now show that the space of bdd. sol. is exactly 1-dimensional.

Thm:  $\phi \in \mathbb{R}^d$ ,  $F: \partial A \rightarrow \mathbb{R}$  bdd., Then the only bdd.  $F: A \rightarrow \mathbb{R}$  s.t.

Sublinear  
F is suff. also  
for this repr.  
in 1.3.

$$(\Delta F)(x) = 0 \quad \forall x \in A, \quad F(x) = F(x) \quad \forall x \in \partial A$$

are of the form  $F(x) = E^x F(S_{\tau_A}^x \mathbb{1}_{\{\tau_A < \infty\}}) + b P^x(\tau_A = \infty)$

for some  $b \in \mathbb{R}$ .

Can also write lazy RW here

Proof: Let  $S_n$  be a lazy SRW. That is  $P(X_1 = 0) = \frac{1}{2}$ ,  $P(X_1 = \pm e_i) = \frac{1}{4}$ .

By the prev. thm, we may assume  $\exists x \in A$  s.t.  $P^x(\tau_A = \infty) > 0$ .

Let  $F$  be a bdd. sol., then  $M_n := F(S_{n \wedge \tau_A}^x)$  is a bdd. mart. Thus

$$F(x) = E^x M_n = E^x F(S_{n \wedge \tau_A}^x) = E^x F(S_n^x) - E^x F(S_n^x) \mathbb{1}_{\{\tau_A < n\}} + E^x F(S_n^x) \mathbb{1}_{\{\tau_A < n\}}$$

for any bdd. cont. of  $F$

Since we have a successful coupling of  $S_n^x$  under  $P^x$  and  $P^y$   $\forall x, y$   
we deduce  $|E^x F(S_n^x) - E^y F(S_n^x)| \leq 2 \|F\|_\infty P(\tau > n) \xrightarrow{n \rightarrow \infty} 0$

Thus  $|F(x) - F(y)| \leq 2 \|F\|_\infty (P^x(\tau_A < \infty) + P^y(\tau_A < \infty))$ .

Exercise: Let  $U_\epsilon = \{x \in A \mid P^x(\tau_A = \infty) \geq 1 - \epsilon\}$ . Exercise:  $U_\epsilon \neq \emptyset \quad \forall \epsilon \in (0, 1)$ .  
since  $\exists x \in A \quad P^x(\tau_A = \infty) > 0$ .

Thus  $\forall x, y \in U_\epsilon$ ,  $|F(x) - F(y)| \leq 4\epsilon \|F\|_\infty$

and hence  $\exists b = \lim_{n \rightarrow \infty} F(x_n)$  for any  $x_n \in U_{1/n}$  s.t.  $|F(x) - b| \leq 4\epsilon \|F\|_\infty$   
 $\forall \epsilon$  and  $x \in U_\epsilon$ .

Exercise: Let  $p_\epsilon = \min\{n \mid S_n^x \in U_\epsilon\}$ . Exercise (related to prev. one)  $P^x(\tau_A \leq p_\epsilon) \leq \epsilon$   
 $\forall \epsilon$  and  $x \in A$ .

Then  $\forall x \in A$ ,  $F(x) = E^x F(S_{p_\epsilon}^x) = E^x F(S_{p_\epsilon}^x) \mathbb{1}_{\{\tau_A > p_\epsilon\}} + E^x F(S_{p_\epsilon}^x) \mathbb{1}_{\{\tau_A \leq p_\epsilon\}}$ .

By bounded conv.  $\lim_{\epsilon \rightarrow 0} E^x F(S_{p_\epsilon}^x) \mathbb{1}_{\{\tau_A \leq p_\epsilon\}} = E^x F(S_{\tau_A}^x) \mathbb{1}_{\{\tau_A < \infty\}}$

(Since  $U_\epsilon \rightarrow \emptyset$  as  $\epsilon \rightarrow 0$  since  $P^x(\tau_A = \infty) < 1 \quad \forall x \in A$ )

Finally  $F(S_{p_\epsilon}^x) \mathbb{1}_{\{\tau_A > p_\epsilon\}} \rightarrow b \mathbb{1}_{\{\tau_A = \infty\}}$  by def. and prop. of  $b$ .

Hence by bdd. conv.  $\lim_{\epsilon \rightarrow 0} E^x F(S_{p_\epsilon}^x) \mathbb{1}_{\{\tau_A > p_\epsilon\}} = b P^x(\tau_A = \infty)$ .

Remark:  $b$  is the boundary data "at infinity". The fact we have only

one such number is the same as saying the Poisson boundary is trivial.  
Repr. like this is the reason for the name Poisson boundary.

Def.: Poisson kernel  $H_A: A \times \partial A \rightarrow [0, 1]$ ,  $H_A(x, y) = P^x(\tau_A < \infty, S_{\tau_A}^x = y)$   
and also  $H_A(x, \infty) = P^x(\tau_A = \infty)$  (slightly abusing notation)  
Then writing  $F(\infty) = b$  we have shown  $F(x) = \sum_{y \in \partial A \cup \{\infty\}} H_A(x, y) F(y)$ .

We turn now to the discrete Poisson equation.

Thm:  $d \neq 1$ .  $A \subseteq \mathbb{Z}^d$ .  $g: A \rightarrow \mathbb{R}$  with finite support. Then the unique bdd. fcn  $f$  on  $\bar{A}$  vanishing on  $\partial A$  and satisfying  $(\mathcal{L}f)(x) = -g(x) \forall x \in A$  is

$$f(x) = \mathbb{E}^x \left( \sum_{j=0}^{\bar{A}-1} g(S_j) \right) = \sum_{y \in A} G_A(x, y) g(y) \text{ with } G_A(x, y) = \mathbb{E}^x \left( \sum_{j=0}^{\bar{A}-1} 1_{\{S_j=y\}} \right)$$

$G_A$  is the Green's fcn for  $A$ .

expected number of visits to  $y$  starting at  $x$  before leaving  $A$ .

Remark: For finite  $A$ ,  $G_A$  is a matrix and defining  $\mathcal{L}_A(x, y) = \begin{cases} -1 & x=y \\ \frac{1}{d} & x \sim y \\ 0 & x \neq y, x \not\sim y \end{cases}$  then the thm. shows  $G_A = -(\mathcal{L}_A)^{-1}$ .

2)  $\mathbb{E}^x(\bar{A}) = \sum_{y \in A} G_A(x, y)$  by taking  $g \equiv 1$ . This is the unique bdd. fcn.  $f: \bar{A} \rightarrow \mathbb{R}$  vanishing on  $\partial A$  and having  $\mathcal{L}f \equiv -1$  on  $A$ .

$\mathcal{L}_A = \begin{pmatrix} -1 & & \\ & \frac{1}{d} & \\ & & \ddots \end{pmatrix}$  in  $d=1$ . On an interval  $A$ .

Proof: It is easy to check this  $f$  satisfies the requirements.

Exercise

For other dir., exercise:  $\forall f: \mathbb{Z}^d \rightarrow \mathbb{R}$ ,  $M_n = f(S_n) - \sum_{j=0}^{n-1} (\mathcal{L}f)(S_j)$  is a martingale.

Let  $f$  be a bdd. sol. We deduce  $M_n = f(S_n) + \sum_{j=0}^{n-1} (\mathcal{L}f)(S_j)$  is a martingale.

Noting  $|M_n| \leq \|f\|_\infty + Y$  with  $Y = \sum_{j=0}^{\bar{A}-1} |g(S_j)|$

We have  $\mathbb{E}^x Y = \sum_y G_A(x, y) |g(y)| < \infty$  and thus by the dom. conv. criterion

for optional sampling:  $f(x) = \mathbb{E}^x M_{\bar{A}} = \mathbb{E}^x \left( \sum_{j=0}^{\bar{A}-1} g(S_j) \right)$  as required.

Remark: To get a sol. not vanishing on  $\partial A$  we add a sol. of the Dirichlet problem.

## 2. Local central limit theorem

Recall that if  $X \sim N(0, n)$  then its density is  $\frac{1}{\sqrt{2\pi n}} e^{-\frac{x^2}{2n}}$ .

We have shown that for 1D SRW, for even  $n+k$  and  $k=0(n^{3/4})$

We have  $P(S_n = k) \sim 2 \cdot \frac{1}{\sqrt{2\pi n}} e^{-k^2/2n}$ . And we could get more explicit error estimates.

Now for large  $n$ ,  $S_n$  is roughly  $N(0, \frac{n}{d} I)$  (since  $\mathbb{E} X_{n+1}^2 = \frac{1}{d}$ )

If  $X \sim N(0, \frac{1}{d} I)$  then its density is  $\left(\frac{1}{\sqrt{2\pi/d}}\right)^d e^{-\frac{d|x|^2}{2n}} = \left(\frac{d}{2\pi n}\right)^{d/2} e^{-\frac{d|x|^2}{2n}}$ .

Thus we may expect that if  $k \in \mathbb{Z}^d$  and  $n$  have the same parity

and  $|k|$  is small enough then  $P(S_n = k) \approx 2 \cdot \left(\frac{d}{2\pi n}\right)^{d/2} e^{-\frac{d|k|^2}{2n}} =: \bar{p}_n(k)$

Let  $E(n, k) = \begin{cases} |p_n(k) - \bar{p}_n(k)| & k \text{ and } n \text{ have same parity} \\ 0 & \text{otherwise} \end{cases}$

Thm. (LCLT):  $\sup_k |E(n, k)| \leq O(n^{-(d+2)/2})$  as  $n \rightarrow \infty$ .

We will not prove this now (possibly later?). There are two main approaches to the proof.

1) Note the vector  $(n_1, \dots, n_d)$  of the number of moves in each coord. has a multinomial dist. Show a local limit thm for  $\mathbb{Z}^d$  and rely on the 1D results in each coord. This is easier in cont. time, when we make a Poisson number of moves since then coords. are ind.

2) Write the Fourier transforms of  $P_n(x)$  and  $\overline{P}_n(k)$  and note that  $\|P_n(k) - \overline{P}_n(k)\|_\infty \leq \| \hat{P}_n(k) - \hat{\overline{P}}_n(k) \|_1$ . Evaluate by Taylor expansion  $\hat{P}_n(k)$  when  $n$  and  $k$  have diff. parity / near origin and by triangle ineq. elsewhere.

Prop. (large dev.):  $\exists C_1, C_2 > 0$  s.t.  $\mathbb{P}(|S_n| \geq r\sqrt{n}) \leq C_1 e^{-C_2 r^2} \forall r, n$

Proof: Follows from 1D estimate we proved:  $\mathbb{P}(S_n \geq k) \leq e^{-k^2/2n}$  in 1D  $\forall k, n_0$

### 3. Green's fcn's.

There are 3 types of Green's fcn's. commonly discussed:

1)  $G(x, y) = \mathbb{E}^x(\# \text{visits to } y) = \sum_{n=0}^{\infty} \mathbb{P}^x(S_n = y)$  defined only for  $d \geq 3$   
 $G(x, y) =: G(y-x)$  since we need transience.  
 For  $d=1, 2$ , the potential kernel  $a(x) = \sum_{n=0}^{\infty} (\mathbb{P}_n(0) - \mathbb{P}_n(x)) =: G(0) - G(x)$  plays a similar role as  $-G(x)$ .

2)  $G(x, y; \lambda)$ : Let  $T$  be a geometric "killing time" with parameter  $1-\lambda$  (chance to be killed at every step).  $\mathbb{P}(T > n) = \lambda^n$ .

Then  $G(x, y; \lambda) = \mathbb{E}^x(\# \text{of visits to } y \text{ before time } T) = \sum_{n=0}^{\infty} \mathbb{P}^x(S_n = y) \lambda^n$ .

3)  $G_A(x, y) = \mathbb{E}^x(\# \text{of visits to } y \text{ before leaving } A)$ .

These are more useful than, say, number of visits to  $y$  by some fixed time  $N$  since they are "Markovian".

After  $n$  steps, we are in same situation for each of them.

Their asymptotics can, in case 1, 2, be deduced from the LGT and large dev. results.

Application: range of 1D and 2D RW: Consider a geometrically

killed SRW as above. Let  $R = R^{\geq 1}$  be the number of distinct positions it visited. Then, by a variant of the RSTW

argument, we have  $\mathbb{E}R = \mathbb{E}T \cdot \mathbb{P}(\forall n \geq 1, S_n \neq 0) = \frac{1}{1-\lambda} \mathbb{P}(\text{no return to } 0)$ .

However, since the number of returns to 0 is geometric,

we have  $\mathbb{P}(\text{no return to } 0) = G(0; \lambda)^{-1}$

Exercise 3

Finally,  $G(0; \lambda) = \sum_{n=0}^{\infty} \lambda^n p_n(0) = \sum_{n=0}^{\infty} \lambda^n (p_n(0) + O(\frac{1}{n^{d+2/2}})) =$   
 $= \sum_{n=0}^{\infty} \lambda^n (\frac{c_d}{n^{d/2}} + O(\frac{1}{n^{d+2/2}})) \approx c_d F(\frac{1}{1-\lambda})$  where  $F(s) = \begin{cases} \sqrt{s} & d=1 \\ \log s & d=2 \end{cases}$

(We have  $c_d = (\frac{d}{2\pi})^{d/2} \cdot \begin{cases} \Gamma(d/2) & d=1 \\ 1 & d=2 \end{cases} = \frac{\sqrt{d}}{c_1}$ )

Thus if  $\lambda = 1 - \frac{1}{n}$  (about  $n$  steps) then  $ER \sim \begin{cases} \frac{n}{c_1 \sqrt{n}} & d=1 \\ \frac{n}{c_2 \log n} & d=2 \end{cases}$

In particular we may deduce the non-trivial result that an  $n$ -step 2D ERW visits  $\frac{n}{c_2 \log n}$  vertices on average (quantitative recurrence) or RSW

Lecture 8 29.7.11 (makeup class) (Scribe)

$c_d e^{-c_d r^2}$

Reminder of LCLT:  $\sup_{\substack{x \in \mathbb{Z}^d \\ n, x \text{ of same parity}}} \frac{E(n, x)}{E(n, 0)} \left| \frac{P(S_n = x) - 2^{d/2} \left(\frac{d}{2\pi n}\right)^{d/2} e^{-d|x|^2/2n}}{E(n, 0)} \right| \leq O(n^{-(d+2/2)}) (n \rightarrow \infty)$   
 Large dev. prop:  $\exists c_d, c_d' > 0, \forall n, P(|S_n| \geq 2\sqrt{dn}) \leq e^{-c_d n}$

Reminder of types of Green's fcn's, correspond to types of transient walks:

- 1) Whole space:  $G(x, y) = E^x(\text{visits to } y) = \sum_{n=0}^{\infty} P^x(S_n = y) = G(x, y)$   
 In  $d=1, 2$ , potential kernel  $a(x) = \sum_{n=0}^{\infty} (P(S_n = 0) - P(S_n = x)) = "G(0) - G(x)"$ .
- 2) Killed walk:  $\tau \sim \exp(1-\lambda)$  ind. of walk.  $G(x) = \sum_{n=0}^{\infty} P(S_n = x, \tau > n)$
- 3) Domain Green's fcn:  $A \subseteq \mathbb{Z}^d, G_A(x, y) = E^x(\# \text{visits to } y \text{ before leaving } A)$   
 $= \sum_{n=0}^{\infty} P^x(S_n = y, \bar{\tau}_A > n)$   
 Lemma:  $G_A(x, y) = G_A(y, x) \forall x, y \in A$   
 Proof: By traversing paths backward  
 $P^x(S_n = y, \bar{\tau}_A > n) = P^y(S_n = x, \bar{\tau}_A > n)$

Asymptotics for Green's fcn's  $|x|$  is Euclidean norm

Thm: For  $d \geq 3$ ,  $G(x) \sim a_d |x|^{2-d}$  as  $x \rightarrow \infty$ , with  $a_d = \frac{d}{2} \pi^{-(\frac{d}{2}-1)} \pi^{-d/2} = \frac{2}{(d-2)\omega_d}$

More precisely  $\forall \alpha < d, G(x) = a_d |x|^{2-d} + o(|x|^{-\alpha})$  ( $\omega_d = \text{Vol}(B(0,1))$  in  $\mathbb{R}^d$ )

Proof: Fix  $x \neq 0$  of even parity. Let  $\bar{p}(2n, x) = 2 \left(\frac{d}{4\pi n}\right)^{d/2} e^{-d|x|^2/4n}$   
 $\sum_{n=1}^{\infty} \bar{p}(2n, x) = \sum_{n=1}^{\infty} 2 \left(\frac{d}{4\pi n}\right)^{d/2} e^{-d|x|^2/4n} \stackrel{\text{exercise}}{\rightarrow} \int_0^{\infty} 2 \left(\frac{d}{4\pi t}\right)^{d/2} e^{-d|x|^2/4t} dt + O(|x|^{-d}) =$   
 $= \frac{d}{2} \pi^{-(\frac{d}{2}-1)} \pi^{-d/2} |x|^{2-d} + O(|x|^{-d}),$  as  $x \rightarrow \infty$ .

Now  $G(x) = \sum_{n=0}^{\infty} P(S_{2n} = x) = \sum_{n=0}^{\infty} \bar{p}(2n, x) + \sum_{n=0}^{\infty} E(2n, x)$

the second.

It remains to show that  $\forall \alpha < d, \sum_{n=0}^{\infty} |E(2n, x)| = o(|x|^{-\alpha})$  as  $x \rightarrow \infty$ .

This follows by considering separately  $n \ll x^2$  and  $n \gg x^2$  and using the large dev. estimate in the first regime and the LCLT in

Remark: In the Continuum, for BM,  $G(x) = a_d |x|^{2-d}$ .

It is a harmonic fn.  $\forall x \neq 0$  and has  $-\Delta$  as Laplacian at 0.

For  $x$  of odd parity we get the result since  $G$  is harmonic off 0, by averaging over the neighbors of  $x$ .

$A_{n,m}$  is better

Remark: It is actually true that the error term is  $O(|x|^{-d})$ .

Application (exiting annuli and transience): For  $n < m$ , let  $A = \{x \mid n < |x| < m\}$

And  $\gamma = \min\{i \mid S_i \notin A\}$ , then  $P^x(S_{\gamma} \leq n) = \frac{|x|^{2-d} - m^{2-d} + O(n^{1-d})}{n^{2-d} - m^{2-d}}$  as  $n \rightarrow \infty$

Note  $P^x(S_{\gamma} \leq n) = \left(\frac{|x|}{n}\right)^{2-d} + O(n^{-1})$  as  $n \rightarrow \infty$ . Quantitative transience.

Proof:  $G(x)$  is harmonic off 0, hence  $M_{\gamma} = G(S_{\gamma})$  is a b.v.l. mart.

By optional sampling  $G(x) = E^x(M_{\gamma}) = P^x(S_{\gamma} \leq n) E^x(G(S_{\gamma}) \mid S_{\gamma} \leq n) + (1 - P^x(S_{\gamma} \leq n)) E^x(G(S_{\gamma}) \mid S_{\gamma} \geq m)$

and the result follows from the prev. thm. using the weaker

estimate  $G(x) = a_d |x|^{2-d} + O(|x|^{1-d})$ .

Similarly, let  $C_n = \{x \mid |x| < n\}$  and  $\eta = \min\{i \mid S_i \notin C_n\}$

then for  $x \in C_n$ ,  $P^x(S_{\eta} = 0) = \frac{a_d}{G(0)} (|x|^{2-d} - n^{2-d}) + O(|x|^{1-d})$  (\*)

To calculate Green's fn. in a domain, we have the following

Prop.: For  $d \geq 3$ , and finite  $A \subset \mathbb{Z}^d$ ,  $\forall x, z \in A$ ,

$$G_A(x, z) = G(z-x) - \sum_{y \in \partial A} H_{\partial A}(x, y) G(z-y)$$

$\partial A = \{y \notin A \mid y \sim z \in A\}$

Proof:  $G_A(x, z) = E^x \left( \sum_{j=0}^{\tau-1} \mathbb{1}_{\{S_j = z\}} \right) = E^x \left( \sum_{j=0}^{\tau-1} \mathbb{1}_{\{S_j = z\}} - \sum_{j=\tau}^{\infty} \mathbb{1}_{\{S_j = z\}} \right)$ .  
 If  $\tau = \min\{i \mid S_i \notin A\}$ ,  $H_{\partial A}(x, y) = P(S_{\tau} = y)$

Cor.:  $G_{C_n}(x, 0) = a_d (|x|^{2-d} - n^{2-d}) + O(|x|^{1-d})$

Proof:  $G_{C_n}(x, 0) = P^x(S_{\eta} = 0) G_{C_n}(0, 0)$ ,  $G_{C_n}(0, 0) = G(0) + O(n^{2-d})$  and (\*)

Dimensions 1, 2: The analogous results for  $d=1, 2$  are:

For  $d=1$ ,  $a(x) = |x|$

For  $d=2$ ,  $\exists k$  s.t.  $\forall x < 2$   $a(x) = \frac{2}{\pi} \ln|x| + k + o(|x|^{-\alpha})$

(and  $k = \frac{2\gamma}{3} + \frac{3}{\pi} \ln 2$ ,  $\gamma = \text{Euler's const.} = \lim_{n \rightarrow \infty} \sum_{j=1}^n \frac{1}{j} - \ln n$ )

$G_A(x, z) = \sum_{y \in \partial A} H_{\partial A}(x, y) a(y-x) - a(z-x)$  for finite  $A$ ,  $x, z \in A$ .

For  $A = \{x \mid n < |x| < m\}$ ,  $\gamma = \min\{i \mid S_i \notin A\}$ ,  $x \in A$ ,

$$P^x(S_{\gamma} \leq n) = \frac{\ln m - \ln|x| + O(n^{-1})}{\ln m - \ln n}$$

Quantitative recurrence

And for  $d=2$   $G_n(x,0) = \frac{2}{\pi} (\ln n - \ln|x|) + o(|x|^{-\alpha}) + O(n^{-1})$ .

(important for <sup>discrete</sup> Gaussian free field, say)

Martin capacity for Markov chains (following Benjamini-Pemantle-Peres 1995)

Def.: Given a measure space  $(\Lambda, \mathcal{F})$ , a measurable function

(called kernel)  $F: \Lambda \times \Lambda \rightarrow [0, \infty]$  and a finite measure  $\nu$  on  $\Lambda$ ,

the F-energy of  $\nu$  is  $I_F(\nu) = \int_{\Lambda} \int_{\Lambda} F(x,y) d\nu(x) d\nu(y)$

The capacity of  $\Lambda$  wrt.  $F$  is  $\text{Cap}_F(\Lambda) = (\inf_{\nu} I_F(\nu))^{-1}$  with  $\infty^{-1} = 0$

where the infimum is over all prob. meas.  $\nu$  on  $\Lambda$ .

For us,  $\Lambda$  will be countable and  $\mathcal{F}$  will be all subsets, but these ideas work equally well in  $\mathbb{R}^d$ .

Furthermore define the asymptotic capacity of  $\Lambda$  wrt.  $F$

by  $\text{Cap}_F^{(\infty)}(\Lambda) := \inf_{\Lambda_0 \text{ finite}} \text{Cap}_F(\Lambda \setminus \Lambda_0)$ . ← note capacity is monotone in the set.

Now let  $\{X_n\}_{n=0}^{\infty}$  be a Markov chain on a set  $Y$  with transition prob.  $P(x,y)$  and define its Green's fcn.:

$$G(x,y) := \sum_{n=0}^{\infty} P^n(X_n=y) = \sum_{n=0}^{\infty} P^{(n)}(x,y).$$

We want to estimate the prob. that  $X_n$  hits a set  $\Lambda \subseteq Y$ .

Thm.: Let  $\{X_n\}$  be a transient Markov chain (no state is visited i.o.)

on a countable state space  $Y$  and  $\Lambda \subseteq Y$ . then  $\forall p \in Y$ :

$$\frac{1}{2} \text{cap}_K(\Lambda) \leq P^p(n \geq 0, X_n \in \Lambda) \leq \text{Cap}_K(\Lambda) \quad (1)$$

$$\frac{1}{2} \text{cap}_K^{(\infty)}(\Lambda) \leq P^p(X_n \in \Lambda \text{ i.o.}) \leq \text{Cap}_K^{(\infty)}(\Lambda) \quad (2)$$

where  $K$  is the Martin kernel:  $K(x,y) := \frac{G(x,y)}{G(p,y)}$  defined using  $p$ .

Remarks: 1) Historically, the Green's fcn. itself is used

as the kernel for SRW or BM. However, note that the Green's fcn. is translation-inv. whereas the hitting prob.

are not. Furthermore, for BM, hitting prob. are scale-inv.

whereas the Green's fcn. is not. The Martin kernel remedies both issues.

2) It suffices instead of transience to assume  $G(x,y) < \infty \forall x,y \in \Lambda$ .

3) The Martin kernel can be replaced by its sym. version  $\frac{1}{2}(K(x,y) + K(y,x))$  without affecting the capacities.

4) The thm. extends to having init. dist.  $\pi$  by adding an abstract state  $p$  which moves according to  $\pi$  on the first step.

Proof: The right hand ineq. follows by taking  $\nu$  to be the harmonic measure. That is, let  $\tau = \min\{n \geq 0 \mid X_n \in \Lambda\}$  and  $\nu(x) = P^p(X_\tau = x, \tau < \infty)$

$\forall y \in \Lambda \int G(x,y) d\nu(x) = \sum_{x \in \Lambda} P^p(X_\tau = x) G(x,y) = G(p,y)$ . if  $\nu(\Lambda) = 0$   
the right hand ineq. is trivial.

Thus  $\int K(x,y) d\nu(x) = 1 \quad \forall y \in \Lambda$ . Consequently:

$$I_K\left(\frac{\nu}{\nu(\Lambda)}\right) = \nu(\Lambda)^{-2} I_K(\nu) = \nu(\Lambda)^{-1} \Rightarrow \text{cap}_K(\Lambda) \geq \nu(\Lambda).$$

For the left hand ineq., we use the second moment method.

Given a prob. measure  $\mu$  on  $\Lambda$ , let

$$Z = \int_{\Lambda} G(p,y)^{-1} \sum_{n=0}^{\infty} \mathbb{1}_{\{X_n=y\}} d\mu(y).$$

Using Fubini,  $E_p Z = 1$ . We bound the second moment

$$\begin{aligned} E_p Z^2 &= E_p \int_{\Lambda} \int_{\Lambda} G(p,y)^{-1} G(p,x)^{-1} \sum_{m,n=0}^{\infty} \mathbb{1}_{\{X_m=x, X_n=y\}} d\mu(x) d\mu(y) \leq \\ &\leq 2 E_p \int_{\Lambda} \int_{\Lambda} G(p,y)^{-1} G(p,x)^{-1} \sum_{0 \leq m \leq n < \infty} \mathbb{1}_{\{X_m=x, X_n=y\}} d\mu(x) d\mu(y) \end{aligned}$$

For each  $m$ , we have  $E_p \sum_{n=m}^{\infty} \mathbb{1}_{\{X_n=x, X_n=y\}} = P_p(X_m=x) G(x,y)$

Therefore,  $E_p Z^2 \leq 2 \int_{\Lambda} \int_{\Lambda} G(p,y)^{-1} G(x,y) d\mu(x) d\mu(y) = 2 I_K(\mu)$

By the second moment ineq. (Cauchy-Schwarz),

$$P_p(\exists n, X_n \in \Lambda) \geq P_p(Z > 0) \geq \frac{(E_p Z)^2}{E_p(Z^2)} \geq \frac{1}{2 I_K(\mu)}$$

Maximizing over  $\mu$  we have  $P_p(\exists n, X_n \in \Lambda) \geq \frac{1}{2} \text{cap}_K(\Lambda)$

To infer (2) from (1), note that since  $\{X_n\}$  is transient,

$$\{X_n \in \Lambda \text{ i.o.}\} \stackrel{(\text{mod } 0)}{=} \bigcap_{\Lambda_0 \text{ finite}} \{\exists n, X_n \in \Lambda \setminus \Lambda_0\}$$

then by monotonicity, if  $\{\Lambda_n\}_{n=0}^{\infty}$  is a seq. of finite sets

increasing to  $\Lambda$ , then  $P(X_n \in \Lambda \text{ i.o.}) = \lim_{n \rightarrow \infty} P(\exists n, X_n \in \Lambda \setminus \Lambda_n)$

$$\text{and } \text{cap}_K^{(\infty)}(\Lambda) = \lim_{n \rightarrow \infty} \text{cap}_K(\Lambda \setminus \Lambda_n)$$

There are many nice corollaries in the paper of BPP. Here is one:

Cor.: Define the Riesz-type kernel  $F_\alpha(x, y) = \frac{1}{1 + \|x - y\|^\alpha}$  on  $\mathbb{Z}^d$

where  $\|\cdot\|$  is any norm. By the prev. thm. and the Hewitt-Savage 0-1 law, SRTW on  $\mathbb{Z}^d$  visits a set  $A$  i.o. a.s. iff  $\text{Cap}_{F_\alpha}^{(\infty)}(A) > 0$ .

Remark: For BM, a Borel set is called polar if  $\forall x$ ,

$P^x(B_t \in A \text{ for some } t > 0) = 0$ . Kakutani (44) showed that  $A$

is polar iff  $\text{Cap}_F(A) = 0$  for  $F(x) = \begin{cases} |\log(|x|)| & d=2 \\ |x|^{2-d} & d \geq 3 \end{cases}$

The prev. thm. (with  $k(x, y) = \frac{1}{|x-y|^{d-2}}$  for  $d \geq 3$ ) can be shown to apply also to BM and to give a quantitative extension to Kakutani's thm.

Put after def. of Capacity as motivation

### Intersection of random walks (following Lawler-Limic)

Let  $S[n_1, n_2] = \{S_n \mid n_1 \leq n \leq n_2\}$ . We have

Thm.:  $\exists c_1, c_2 > 0, \forall n \geq 2, c_1 \phi(n) \leq P(S[0, n] \cap S[2n, 3n] \neq \emptyset) \leq c_2 \phi(n)$

With  $\phi(n) = \begin{cases} 1 & d < 4 \\ 1/\log n & d = 4 \\ 1/(d-4)^2 & d > 4 \end{cases}$

Remarks: 1) Same holds for any sym., fin. supp. <sup>irred.</sup> walk.

2) Taking  $n \rightarrow \infty$  we get  $P(B[0, 1] \cap B[2, 3] \neq \emptyset) \begin{cases} > 0 & d \leq 3 \\ = 0 & d \geq 4 \end{cases}$  for a BM  $B$ .

Proof: (Following Lawler-Limic) Trivial UB for  $d \leq 3$ . LB for  $d \leq 2$

Follows from  $d=3$ . Hence assume  $d \geq 3$ . We use the second

moment method. Let  $J_n = \sum_{j=0}^n \sum_{k=2n}^{3n} \mathbb{1}_{(S_j = S_k)}$   $K_n = \sum_{j=0}^n \sum_{k=2n}^{3n} \mathbb{1}_{(S_j \neq S_k)}$

Using the LGLT and large dev.  $c_1 n^{\frac{4-d}{2}} \leq \mathbb{E} J_n \leq \mathbb{E} K_n \leq c_2 n^{\frac{4-d}{2}}$

Similarly, the LGLT and large dev. give  $\mathbb{E} J_n^2 \leq \begin{cases} cn & d=3 \\ c \log n & d=4 \\ cn^{(4-d)/2} & d \geq 5 \end{cases}$

since  $P(K_n \geq 1) \leq \mathbb{E} K_n$  and  $P(J_n > 0) \geq \frac{\mathbb{E} J_n^2}{\mathbb{E} J_n}$

the lower bounds for all  $d$  and the upper bounds for all  $d \neq 4$  follow.

The upper bound for  $d=4$  needs more work (estimating  $\mathbb{E}(K_n \mid K_n \geq 1) \geq c \log n$ )

and we will not do it here (see Lawler-Limic).

Consider explaining easy example in Lawler-Limic after proof.

like expected number of intersections for two walks starting at origin



Construction of BM: Give some prop. of Gaussian RVs.

Def.: We say that  $X = (X_1, \dots, X_n)$  is dist. as a Gaussian vector if  $X \stackrel{d}{=} AY + b$  for  $Y = (Y_1, \dots, Y_k) \text{ IID } N(0, 1)$  and some  $k \geq 1$ ,  $A \in M_{n \times k}$  and  $b \in \mathbb{R}^n$ .

Thm.: The dist. of a Gaussian vector is uniq. determined by its mean vector and cov. matrix. (proof, e.g., by Fourier transform)

Cor.: If  $X_1, X_2$  are <sup>dist.</sup> ind.  $N(0, 1)$  then  $\frac{X_1 + X_2}{\sqrt{2}}, \frac{X_1 - X_2}{\sqrt{2}}$  are also ind.  $N(0, 1)$ .

Prop.: If  $\{X_n\}_{n \geq 1}$  is a seq. of Gaussian vectors in  $\mathbb{R}^n$  conv. in dist. to  $X$  then  $X$  is  $N(\mu, \Sigma)$  with  $\mu$  and  $\Sigma$  the limits (which exist) of the mean and cov. of  $X_n$ .

Prop. (Tail estimate): If  $X \sim N(0, 1)$  then

$$\frac{x}{x^2 + 1} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \leq \mathbb{P}(X > x) \leq \frac{1}{x} \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

Did existence thm. and showed upper and lower bounds for modulus of continuity.

Lecture 10 13.5.11 (make-up class)

Reminders: BM is a random cont. fcn. on  $[0, \infty)$  starting at  $x$

s.t.  $B(0) = x$  a.s., independent incr. and  $B(t+h) - B(t) \sim N(0, h)$ .

Discussed modulus of cont. of  $B[0, 1]$ :

- 1)  $\exists c > 0$  s.t. a.s.  $\forall$  suff. small  $h$  and all  $0 \leq t \leq t-h$  we have  $|B(t+h) - B(t)| \leq c\sqrt{h \log(1/h)}$
- 2)  $\forall c < \sqrt{2}$ , a.s.  $\forall \epsilon > 0 \exists 0 < h < \epsilon$  and  $t \in [0, t-h]$  s.t.  $|B(t+h) - B(t)| \geq c\sqrt{h \log(1/h)}$ .

Hölder cont.: Def.:  $f: [0, \infty) \rightarrow \mathbb{R}$  is said to be locally  $\alpha$ -Hölder cont. at  $x \geq 0$  if  $\exists \epsilon > 0, c > 0$  s.t.  $|f(x) - f(y)| \leq c|x - y|^\alpha \forall y \geq 0$  with  $|x - y| < \epsilon$ .

$\alpha > 0$  is called the Hölder exp. and  $c > 0$  the Hölder const.

Cor.:  $\forall x \in \mathbb{Z}$ , a.s. BM is locally  $\alpha$ -Hölder cont. at every point (simultaneously).

Proof: It follows immed. from the mod. of cont. results for  $B[0, 1]$ . Hence it is also true  $\forall k$  for  $B_k(t) := B(t+k) - B(k)$  for  $t \in [0, 1]$ .

o p. 15, why 2. is bound on  $\mathbb{R}^n$ ?  
 o p. 21  $k=2$  -12 why restrict  $\leq 2$   $k$ ?  
 o Proof of thm. 1.35?  
 o in proof of thm. 1.30. do need for  $2^k$  interv. and BC.  
 o p. 11, since def. of  $F_n$  has  $z + (1/2 - \ln n)/n$ , no need for  $\sqrt{2 \log 2}$ , can take any  $c > \sqrt{2 \log 2}$ .  
 o p. 15, can also bound  $|F_n(t+h) - F_n(t)|$  by  $\|F_n\|_\infty$  without 2 by special def. of  $F_n$ .  
 o Proof of Wiener thm. 1.30. Would help to elaborate more on why  $B$  had ind. incr. on  $D_n$ . Maybe better to start with the  $F_n$ .  
 o In def. of BM, what is the  $r$ -alg. for its space? Would help to understand whether its dist. is uniq. specified.

Remark: It is also true that for any  $\alpha > \frac{1}{2}$ , a.s. at every point

BM is not loc.  $\alpha$ -Hölder cont. This may be an exercise.

Exercise

Points where BM is  $\frac{1}{2}$ -Hölder cont. exist, but are very rare (See martingales - zeros under "slow points", we will not prove this)

Non-differentiability: Thm. (Paley, Wiener, Zygmund 1933): A.s. BM is

not diff. anywhere. moreover,  $\forall t$ , either  $D^*B(t) = \infty$

or  $D_*B(t) = -\infty$  or both, where  $D^*f(t) = \limsup_{h \downarrow 0} \frac{f(t+h) - f(t)}{h}$   
 $D_*f(t) = \liminf_{h \downarrow 0} \frac{f(t+h) - f(t)}{h}$

Proof: It suffices to prove the thm. for all  $t \in [0, 1]$ .

If there exists  $t_0 \in [0, 1]$  with  $-\infty < D_*B(t) \leq D^*B(t) < \infty$

then  $\limsup_{h \downarrow 0} \frac{|B(t+h) - B(t)|}{h} < \infty$  and hence  $\exists M, t_0 \in [0, 1]$

s.t.  $\sup_{h \in [0, 1]} \frac{|B(t_0+h) - B(t_0)|}{h} \leq M$ . It suffices to show this

has zero prob. for every  $M$ .

Fix  $M$  and  $n \in \mathbb{N}$  and suppose  $t_0 \in [\frac{k-1}{n}, \frac{k}{n}]$ . It follows that

for all  $1 \leq j \leq n-1$  we have

$$|B(\frac{k+j}{n}) - B(\frac{k-1+j}{n})| \leq |B(\frac{k+j}{n}) - B(t_0)| + |B(\frac{k-1+j}{n}) - B(t_0)| \leq M \cdot \frac{2j+1}{n}$$

Let  $\mathcal{A}_{nk} := \{ |B(\frac{k+j}{n}) - B(\frac{k-1+j}{n})| \leq M \frac{2j+1}{n} \text{ for } j=1, 2, 3, \dots \}$

Then by ind. of the increments:  $P(\mathcal{A}_{nk}) = \prod_{j=1}^{\infty} P(|B(\frac{k+j}{n}) - B(\frac{k-1+j}{n})| \leq M \frac{2j+1}{n}) = \prod_{j=1}^{\infty} P(|Z| \leq \frac{7M}{\sqrt{n}}) = \prod_{j=1}^{\infty} \frac{2}{\sqrt{n}}$

$Z \sim N(0, 1)$

since the normal density is odd by  $\frac{1}{2}$ . Hence

$$P(\bigcup_{k=1}^n \mathcal{A}_{nk}) \leq n \left( \frac{2}{\sqrt{n}} \right)^3 \xrightarrow{n \rightarrow \infty} 0$$

Since existence of  $M$  and  $t_0$  as above implies that  $\bigcup_{k=1}^n \mathcal{A}_{nk}$  holds for all  $n \geq n_0$ , the thm. follows.

Dist. Prop. of the BM process

Scaling inv.: If  $B$  is a standard BM and  $a > 0$  the process

$\{X(t) | t \geq 0\}$ , defined by  $X(t) := \frac{1}{a} B(a^2 t)$  is also a standard BM.

Proof: Need only to check that increments have correct dist.

Remark: Individual paths change, but the overall dist. does not.

Application: For  $a < 0 < b$ , let  $T(a,b) := \min\{t \geq 0 \mid B(t) \in [a,b]\}$ . (letter  $a, b > 0$  and  $\in [-a, b]$ ?)

Then letting  $X(t) := \frac{1}{|a|} B(a^2 t)$  we have  $\mathbb{P}(T(a,b) = a^2 \min\{t \geq 0 \mid X(t) \in [-1, \frac{b}{|a|}]\}) =$

In particular  $\mathbb{P}(T(-b,b) = c b^2)$  for some  $c$ .  $a^2 \mathbb{P}(T(-1, \frac{b}{|a|}))$

Similarly  $\mathbb{P}(B(T(a,b)) = a) = \mathbb{P}(X(T(-1, \frac{b}{|a|})) = -1) = f(\frac{b}{|a|})$  only.

Time inversion: If  $B$  is a standard BM then  $\{X(t) \mid t \geq 0\}$  defined

by  $X(t) = \begin{cases} t B(1/t) & t > 0 \\ 0 & t = 0 \end{cases}$  is also a standard BM.

Proof: Easy to check finite dim. dist. remain the same by checking cov structure. It remains only to show cont. at 0.

Since dist. on rationals is same as for BM, a.s.  $\lim_{t \downarrow 0} X(t) = 0$ .

By cont. of  $X$  on  $(0, \infty)$  we get cont. at 0.

Remark: The Ornstein-Uhlenbeck process is defined by  $Y(t) := e^{-t} B(e^{2t})$  for  $t \in \mathbb{R}$ .

It is a stat. Markov process which is a diffusion

with drift towards origin prop. to dist. from it.

One may view time inversion as the statement that  $Y(t)$  is time-reversible.

Relating Prop. at  $\infty$  to Prop. at 0.

Applications: 1) LLN: a.s.,  $\lim_{t \rightarrow \infty} \frac{B(t)}{t} = 0$ . (standardized)

Proof: For  $X(t)$  the time-inv. of  $B$ , we have  $\lim_{t \rightarrow \infty} \frac{B(t)}{t} = \lim_{t \rightarrow \infty} X(\frac{1}{t}) = 0$  a.s.

2) a.s.,  $\limsup_{n \rightarrow \infty} \frac{B(n)}{\sqrt{n}} = \infty$  and  $\liminf_{n \rightarrow \infty} \frac{B(n)}{\sqrt{n}} = -\infty$ . (Remark correct gauge fcn. is the law of the iter. log.)

Proof:  $\forall M, \mathbb{P}(B(n) > M\sqrt{n} \mid \mathcal{F}_n) \in (0, 1/2)$  since  $B(n) = \sum_{i=1}^n (B(i) - B(i-1))$  and the

Hewitt-Savage 0-1 law. Fix  $M$ . Since for any seq. of events

we have  $\mathbb{P}(A_n \mid \mathcal{F}_n) \geq \limsup_{n \rightarrow \infty} \mathbb{P}(A_n)$  (Fatou, since  $\{A_n \mid \mathcal{F}_n\} = \bigcap_{n=1}^{\infty} \bigcup_{k=n}^{\infty} A_k$  and the mon. cont. thm.)

then  $\mathbb{P}(B(n) > M\sqrt{n} \mid \mathcal{F}_n) \geq \limsup_{n \rightarrow \infty} \mathbb{P}(B(n) > M\sqrt{n}) = \mathbb{P}(B(1) > M) > 0$ .

Hence  $\forall M, \mathbb{P}(B(n) > M\sqrt{n} \mid \mathcal{F}_n) = 1$ , proving the Prop.

3) a.s.,  $\limsup_{h \downarrow 0} \frac{B(h)}{\sqrt{h}} = \infty$  and  $\liminf_{h \downarrow 0} \frac{B(h)}{\sqrt{h}} = -\infty$ .

Proof: Let  $X(t)$  be the time inv. process, then  $\limsup_{h \downarrow 0} \frac{B(h)}{\sqrt{h}} = \limsup_{h \downarrow 0} \sqrt{h} X(\frac{1}{h}) =$

4) Letting  $\tau = \inf\{t > 0 \mid B(t) > 0\}$ ,  $\sigma = \inf\{t > 0 \mid B(t) < 0\}$

we have  $\mathbb{P}(\tau = 0) = \mathbb{P}(\sigma = 0) = 1$ .

Proof: Follows from 3.

5) BM has no interval of mono.

Proof: Let  $a < b$  be rational. If  $[a,b]$  is an int. of mono. then the int. of any subinterval has the same sign.

## MARKOV PROP. and Blumenthal's 0-1 law

DEF: J-Jim. BM,  $P^x$  and  $E^x$ .

DEF: Ind. of Proc. (for us) is ind. of Fin. dim. dist.

Thm. (Markov Prop.)

DEF: Filtration, filtered prob. space and adapted.

Relevant filtrations:  $\mathcal{F}^0(t) = \sigma(B(s) | 0 \leq s \leq t)$ ,  $\mathcal{F}^+(t) = \bigcap_{s > t} \mathcal{F}^0(s)$

Thm:  $\{B(t+s) - B(s) | t \geq 0\}$  ind. of  $\mathcal{F}^+(t)$ .

Proof: Approx. points from right. Ind. passes to limit.

Thm. (Blumenthal's 0-1 law): Germ  $\sigma$ -alg.  $\mathcal{F}^+(0)$  is trivial.

Application: Tail  $\sigma$ -alg.  $\mathcal{F} = \bigcap_{t \geq 0} \sigma(B(s) | s \geq t)$  is trivial.

Proof: Time inv. and Blum. 0-1 law.

(Exercise) Remark: In fact,  $\forall A \in \mathcal{F}$ ,  $P^x(A) \in \{0, 1\}$  and this prob. is ind. of  $x \in \mathbb{R}^d$ .

(Reached here!) However, this is not true for the germ field  $\mathcal{F}^+(0)$ .

Lecture 11 16.5.11 (Scribe) (HW) Application of Markov prop. to local maxima thm. (p. 39).

## Strong Markov prop. and reflection principle

DEF: Stopping time in a filtered prob. space.

(minimum of two stop. times) Examples: deterministic, increasing limit, discretization.

Which filtration to use?  $\mathcal{F}^0(t)$  or  $\mathcal{F}^+(t)$ ?  $\mathcal{F}^0(t) \subseteq \mathcal{F}^+(t)$ , hitting a closed set, hitting an open set.

Right Continuity and alternative def. of stopping time.

Introduce  $\mathcal{F}^+(\tau)$ . Strong Markov Prop.

(Skin proof) Reflection principle. The area of planar BM. Cor: <sup>planar</sup> BM does not visit points. zero set of BM is a perfect set.

part. in cont. time, Wald's lemmas for BM, Gambler's ruin for BM.

Skorohod embedding - Azema-Yor embedding, Donsker's inv. principle.

(Remind Markov Prop.) Thm. (Application of Markov prop. to local maxima): For a 1D BM, a.s.,

- 1) Every local max. is strict.
- 2) The set of times of local maxima is countable and dense.
- 3) The global max. is attained at a unique time.



Note that  $\mathcal{F}^+(t)$  is right-cont. in the sense that  $\forall t, \mathcal{F}^+(t) = \bigcap_{\epsilon > 0} \mathcal{F}^+(t+\epsilon)$ .

Lemma: If  $\tau$  is a RV taking values in  $[0, \infty]$  and satisfying

$$\{\tau < t\} \in \mathcal{F}(t) \quad \forall t \text{ wrt. a right-cont. filt. } \{\mathcal{F}(t)\}_{t \geq 0}$$

$\tau$  is a stop. time wrt.  $\{\mathcal{F}(t)\}$ .

Proof:  $\{\tau \leq t\} = \bigcap_{k=1}^{\infty} \{\tau < t + \frac{1}{k}\} \in \bigcap_{k=1}^{\infty} \mathcal{F}(t + \frac{1}{k}) = \mathcal{F}(t)$ . right cont.

DEF: For a stop. time  $\tau$ ,  $\mathcal{F}^+(\tau) = \{A \in \mathcal{A} \mid A \cap \{\tau \leq t\} \in \mathcal{F}^+(t) \quad \forall t\}$ .

All the events known by time  $\tau$ . Note, e.g.,  $\{B(t) \mid t \leq \tau\} \in \mathcal{F}^+(\tau)$

and note we can write  $\{\tau < t\}$  in the above def. by right-cont.

Thm (strong Markov prop) For every a.s. finite stop. time

the process  $\{B(t+\tau) - B(\tau) \mid t \geq 0\}$  is a standard BM ind. of  $\mathcal{F}^+(\tau)$ .

Remark: There are "bad" times too, say  $\tau = \max\{t \leq 1 \mid B(t) = 0\}$ .

Proof: following BM book (p. 73).

Applications: REFLECTION PRINCIPLE: If  $\tau$  is a stop. time

and  $B$  a standard BM then  $B^*$ , called BM reflected at  $\tau$

$$\text{and defined by } B^*(t) := B(t) \mathbb{1}_{\{t \leq \tau\}} + (2B(\tau) - B(t)) \mathbb{1}_{\{t > \tau\}}$$

is also a standard BM.

Proof: For finite  $\tau$ , by the strong Markov prop, both

$$\{B(t+\tau) - B(\tau) \mid t \geq 0\} \text{ and } \{- (B(t+\tau) - B(\tau)) \mid t \geq 0\}$$

are standard BM ind. of  $\mathcal{F}(t) \mid 0 \leq t \leq \tau$ .

Gluing the first to the end of  $\{B(t) \mid 0 \leq t \leq \tau\}$  gives  $B$ .

Gluing the second gives  $B^*$ .

If  $\tau$  is not necessarily finite, we can proceed by applying this to  $\tau \wedge n$  and taking a limit  $n \rightarrow \infty$ . The limit is a.s. cont. and has the correct fin. dim. dist.

Cor: If  $a > 0$  then  $P(M(t) \geq a) = 2P(B(t) \geq a) = P(|B(t)| \geq a)$

PROOF: APPLY REF. PRINCIPLE to first hitting time of  $a$ .

$$\{M(t) \geq a\} = \{B(t) \geq a\} \cup \{B(t) < a, M(t) \geq a\} \\ = \{B^*(t) \geq a\}$$

P. 198, Ex. 5.10 All assume that  $|X| < \infty$ .

P. 45, proof of thm. 2.21, calc.  $P(M(t) \geq a)$  instead since it is only true mod 0 that

$$\{M(t) > a, B(t) < a\} = \{B^*(t) \geq a\}$$

P. 46, proof of thm. 2.24 seems it is not eq. in first display of proof and needs to be  $B$ . As in write up.

P. 47,  $B_0$  not d. BM by def. Replace  $\tau_0 = 0$  a.s. by a.e.

P. 127, thm. 5.15, natural filt. of BM,  $\mathcal{F}^0$  or  $\mathcal{F}^+$ ?

P. 57 In the optional stop. prop. 2.72, can  $\tau$  be infinite? related to 501 ex. 2.6.

(Lévy 1940)  
 2) Area of Planar BM: The range of 2D BM has zero area, a.s.

Remark: BUT it has Hausdorff dim. 2. Much stronger,  $\forall A \subset [0, \infty)$  closed  
 $\dim B(A) = (2 \dim A) \wedge d$ . see McKean thm. 7.33 and Kaufman's thm. 9.28.

Def.:  $\mathcal{L}_d =$  Leb. meas. on  $\mathbb{R}^d$ .  $f * g =$  the conv. of  $f$  and  $g$ ,  $(f * g)(x) = \int f(y)g(x-y)dy$ .

$$A+x := \{a+x \mid a \in A\}$$

Lemma: If  $A_1, A_2$  are Borel sets in  $\mathbb{R}^d$  with pos. meas. then

$$\mathcal{L}_d(\{x \in \mathbb{R}^d \mid \mathcal{L}_d(A_1 \cap (A_2+x)) > 0\}) > 0$$

Proof: Assume wlog  $A_1, A_2$  bdd. Note  $\mathcal{L}_d(A_1 \cap (A_2+x)) = \mathbb{1}_{A_1} * \mathbb{1}_{-A_2}(x)$ .

$$\begin{aligned} \text{By Fubini, } \int_{\mathbb{R}^d} (\mathbb{1}_{A_1} * \mathbb{1}_{-A_2})(x) dx &= \int_{\mathbb{R}^d} \int_{\mathbb{R}^d} \mathbb{1}_{A_1}(w) \mathbb{1}_{-A_2}(x-w) dw dx \\ &= \int_{\mathbb{R}^d} \mathbb{1}_{A_1}(w) \int_{\mathbb{R}^d} \mathbb{1}_{-A_2}(x-w) dx dw = \mathcal{L}_d(A_1) \mathcal{L}_d(A_2) > 0 \end{aligned}$$

Proof of Lévy's thm: It is suff. to show that a.s.  $\mathcal{L}_2(B[0,1]) = 0$ .

Let  $X = \mathcal{L}_2(B[0,1])$ . First we check  $E X < \infty$ . This follows from a tail estimate on the max. of a 1D BM  $W$  since

$$P(X > a) \leq 2 P(\max_{t \in [0,1]} |W(t)| > \frac{\sqrt{a}}{2}) \stackrel{\text{prev. cor.}}{\leq} 8 P(W(1) > \frac{\sqrt{a}}{2}) \leq 8e^{-a/8}$$

Second, by scale inv.  $\{B(t)\}_t$  and  $\{\sqrt{3}B(\frac{t}{3})\}_t$  have the same dist.

$$\text{Hence } E \mathcal{L}_2(B[0,3]) = 3 E \mathcal{L}_2(B[0,1]) = 3 E X.$$

However,  $\mathcal{L}_2(B[0,3]) \leq \sum_{j=0}^2 \mathcal{L}_2(B[j, j+1])$  with equality iff  $\mathcal{L}_2(B[i, i+1] \cap B[j, j+1]) = 0$

$$\text{Moreover } 3 E X = E \mathcal{L}_2(B[0,3]) \leq \sum_{j=0}^2 E \mathcal{L}_2(B[j, j+1]) = 3 E X.$$

Hence a.s.  $\mathcal{L}_2(B[i, i+1] \cap B[j, j+1]) = 0 \quad \forall 0 \leq i < j \leq 2$ .

Now consider  $\{B(t) \mid t \in [0,1]\}$ ,  $Y := B(2) - B(1)$  and  $\{B(t) \mid t \in [0,1]\}$  defined

by  $B'(t) := B(t+2) - Y$ . By the Markov prop.,  $Y$  is ind. of both these processes since  $B'(t) = [B(t+2) - B(2)] + B(1)$ .

Let  $R(x) := \mathcal{L}_2(B[0,1] \cap (x + B'[0,1]))$  and note  $0 = E \mathcal{L}_2(B[0,1] \cap B[2,3]) = E R(Y)$ .

Since  $Y$  is Gaussian, it follows  $R(x) = 0$  for a.e.  $x$ . By prev. lemma

we obtain that a.s.  $\mathcal{L}_2(B[0,1]) = 0$  or  $\mathcal{L}_2(B[2,3]) = 0$ . Since these are ind., we are done.

Cor.: For any  $d \geq 2$  and any  $x, y \in \mathbb{R}^d$   $P_x(Y \in B[0,1]) = 0$  ( $d$ -dim. BM does not hit points for  $d \geq 2$ )

Proof: By Prop., we may assume  $d=2$ . Note  $\int_{\mathbb{R}^2} P_y(x \in B[0,1]) dx = E_y \mathcal{L}_2(B[0,1]) = 0$   
 Hence for every  $y$  and a.e.  $x$ ,  $P_y(x \in B[0,1]) = 0$ . Note Fubini Lévy

$$P_y(x \in B[0,1]) = P_0(x - y \in B[0,1]) = P_0(y - x \in B[0,1]) = P_x(y \in B[0,1]).$$

Finally, to get every  $x$ , note  $P_x(Y \in B[0,1]) = \lim_{\epsilon \downarrow 0} P_x(Y \in B[\epsilon, 1]) = \lim_{\epsilon \downarrow 0} E_x P_{B(\epsilon)}(Y \in [0, \epsilon]) = 0$ .

3) zero set of 1D BM is a perfect set: Proof from BM book, p. 48

Reached here

Using strong Markov prop.

Lecture 12.23.5.11 (scribe) (HW)  
DEF. of martingale, submart., in cont. time.

EX. 1D BM is a martingale.

Prop. (optional stop. thm.) state but do not prove. proof

by approx., see BM book, p. 57.

Prop. (Wald's lemma).

Cor.:  $S \leq T$  stop. times with  $E T < \infty$ :

$$E(B(T)^2) = E(B(S)^2) + E(B(T) - B(S))^2$$

$B(t)^2 - t^2$  is a mart. for BM.

Second Wald's lemma.

Exit prob. and time for a 1D BM.

Skorohod embedding.

DEF: A real-valued stochastic process  $\{X(t) | t \geq 0\}$  is a martingale with respect to a filtration  $\{\mathcal{F}(t) | t \geq 0\}$

if it is adapted to  $\{\mathcal{F}(t)\}_{t \geq 0}$ ,  $E|X(t)| < \infty \forall t \geq 0$

and for any  $0 \leq s \leq t$ , we have  $E(X(t) | \mathcal{F}(s)) = X(s)$  a.s.

Submartingale:  $\geq$ , Supermartingale:  $\leq$

Example: 1D BM is a martingale wrt.  $\mathcal{F}^+(t)$ . For  $s \leq t$ ,

$$E(B(t) | \mathcal{F}^+(s)) = E(B(t) - B(s) | \mathcal{F}^+(s)) + B(s) \stackrel{\text{Markov Prop. wrt. } \mathcal{F}^+(s)}{=} E(B(t) - B(s)) + B(s) = B(s)$$

DEF: The martingale  $\{X(t)\}_{t \geq 0}$  is called cont. if a.s. its sample paths are cont.

Thm. (Optional stopping): If  $\{X(t)\}_{t \geq 0}$  is a cont. mart.

and  $T$  is a stopping time s.t. there exists an int.  $X$

satisfying  $|X(t \wedge T)| \leq X$  a.s. for every  $t$ , then  $E X(T) = X(0)$ .

Proof: Use approx. by a discrete-time mart. and corresponding result in discrete time. Left as exercise.

Exercise

Remark: It is also true that if we have two stopping times  $S \leq T$  a.s. and  $|X(t \wedge T)| \leq X$  for an int.  $X$

then  $E(X(T) | \mathcal{F}(S)) = X(S)$  a.s. This can again be proved using the corresp. discrete-time result.

Prop. (Wald's lemma for BM): If  $T$  is a stopping time for 1D BM (wrt.  $\mathcal{F}^t$ ) and  $E T < \infty$  then  $E B(T) = 0$ .

Remark: In fact, this is still true under the assump.  $E T^2 < \infty$ .

Note it is not true for hitting time  $T_1$  of 1 though  $E T_1^2 < \infty$   $\forall 0 < \epsilon < \frac{1}{2}$

Proof: Let  $M_k := \max_{0 \leq t \leq 1} |B(t+k) - B(t)|$  and  $M = \sum_{k=1}^{\lceil T \rceil} M_k$ .

Note  $\forall t, |B(t \wedge T)| \leq M$  a.s. so if  $E M < \infty$  the Prop. will follow from the optional stop. thm. We have

$$E M = \sum_{k=1}^{\infty} E M_k \mathbb{1}_{\{T > k-1\}} \stackrel{\text{Markov Prop. and } \{T > k-1\} \in \mathcal{F}^+(k-1)}{=} \sum_{k=1}^{\infty} E M_k P(T > k-1) = E M_0 \cdot E(T+1)$$

and  $E M_0 < \infty$  since  $P(M_0 > a) = P(|B(1)| > a)$ .

Lemma: For a 1D BM,  $(B(t)^2 - t)_{t \geq 0}$  is a mart. (wrt.  $\mathcal{F}^t$ ).

Proof: Adaptedness and int. are clear. For all  $s \leq t$ :

$$\begin{aligned} E(B(t)^2 - t | \mathcal{F}^+(s)) &= E((B(t) - B(s))^2 | \mathcal{F}^+(s)) + \underbrace{2E(B(s)B(t) | \mathcal{F}^+(s))}_{=0} - B(s)^2 - t = \\ \text{Markov Prop.} \Rightarrow &= t - s - B(s)^2 - t = B(s)^2 - s. \end{aligned}$$

Prop. (Wald's second lemma for BM): If  $T$  is a stop. time for 1D BM (wrt.  $\mathcal{F}^t$ ) and  $E T < \infty$  then  $E B(T)^2 = E T$ .

Proof: Define  $H_n := \min\{t \geq 0 | |B(t)| = n\}$  and  $T_n := T \wedge H_n$ .

Note  $|B(t \wedge T_n)|^2 - t \wedge T_n \leq n^2 + T$  a.s. Hence by optional stop.

$$E B(T_n)^2 = E T_n. \text{ By Fatou's lemma } E B(T)^2 \leq \liminf_{n \rightarrow \infty} E B(T_n)^2 = \liminf_{n \rightarrow \infty} E T_n \leq E T.$$

For the other dir., note first that

$$E B(T)^2 = E (B(T) - B(T_n))^2 + E B(T_n)^2 + 2E B(T_n)(B(T) - B(T_n)) \geq E B(T_n)^2$$

$$\text{Since } E B(T_n)(B(T) - B(T_n)) = E B(T_n) E (B(T) - B(T_n) | \mathcal{F}^+(T_n)) = 0$$

by the strong Markov prop. and Wald's lemma ( $E(T - T_n | \mathcal{F}^+(T_n)) < \infty$  a.s.)

$$\text{Hence } E B(T)^2 \geq \lim_{n \rightarrow \infty} E B(T_n)^2 = \lim_{n \rightarrow \infty} E T_n = E T. \text{ mono. Contr.}$$

Cor. (Hitting Prob. and time): For  $a, b > 0$ , let  $T = \min\{t \geq 0 | B(t) \in \{-a, b\}\}$ . Then

$$1) P(B(T) = a) = \frac{b}{a+b}, P(B(T) = b) = \frac{a}{a+b}.$$

$$2) E T = ab.$$

Proof: Same as for SRTW using Wald's lemmas.

Skorohod embedding

Moving towards the Donsker Inv. Principle, we seek first to solve the Skorohod embed. Problem. Given a RV  $X$ ,

we seek <sup>an integrable</sup> stopping time  $T$  for BM s.t.  $B(T) \sim X$

By Wald's lemmas,  $E B(T) = 0$  and  $E B(T)^2 = ET$ , hence

this can only be possible for centered  $X$  with finite var.

Thm.: IF  $X$  satisfies  $EX=0, EX^2 < \infty$  then there exists an int. stop. time  $T$  for 1D BM wrt.  $\mathcal{F}^+$  s.t.  $B(T) \sim X$

Example: IF  $X$  takes only two values  $-a, b$  then <sup>(and  $ET = EX^2$ )</sup>

$\min\{t \geq 0 \mid B(t) \in \{-a, b\}\}$  has the required prop. (in particular for  $a=b=1$ )

There are two well-known proofs of the prev. thm.: the Dubins embedding thm. and the Azema-Yor embed. thm. We will present the Azema-Yor embed. See Durrett or Morters-Perez for the other approach.

Thm. (Azema-Yor embed.): IF  $X$  satisfies  $EX=0, EX^2 < \infty$ ,

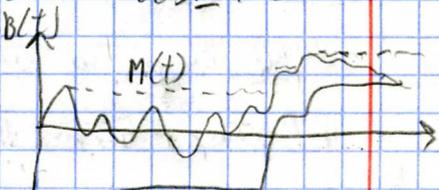
$$\text{let } \psi(x) = \begin{cases} EX \mid X \geq x & \text{if } P(X \geq x) > 0 \\ 0 & \text{otherwise} \end{cases}$$

For a 1D BM  $B$ , let  $M$  be its max. proc.  $M(t) = \max_{0 \leq s \leq t} B(s)$

and define a stop. time  $T$  by  $T := \inf\{t \geq 0 \mid M(t) \geq \psi(B(t))\}$

then  $ET = EX^2$  and  $B(T) \sim X$ .

Proof: We start by proving the thm. when  $X$  is fin. supp.



Lemma: Suppose  $X$  takes only the values

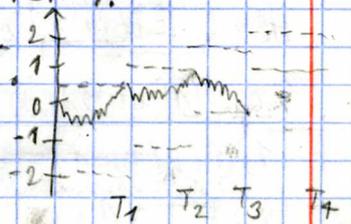
$x_1 < x_2 < \dots < x_n$  and  $EX=0$ . Define

$y_1 < \dots < y_{n-1}$  by  $y_i = \psi(x_{i+1})$  and define stop. times  $T_0 = 0$

and  $T_i = \inf\{t \geq T_{i-1} \mid B(t) \in (x_i, y_i)\}$  for  $i \leq n-1$ .

Then  $T_{n-1}$  satisfies  $ET_{n-1} = EX^2$  and  $B(T_{n-1}) \sim X$ .

$$\psi^{-1}(y_i) = \sup\{b \mid \psi(b) \leq x_i\}$$



Uniform dist. on  $\{-2, -1, 0, 1, 2\}$   
Here  $T = T_3 = T_4$  and  $B(T) = 0$ .

Proof: First note  $y_i \geq x_{i+1}$  with equality iff  $i = n-1$ .

We have  $E T_{n-1} < \infty$  and hence  $E B(T_{n-1})^2 = E T_{n-1}$ . For  $i = 1, \dots, n-1$

$$\text{let } Y_i = \begin{cases} E(X | X \geq x_{i+1}) & \text{if } X \geq x_{i+1} \\ X & \text{if } X \leq x_i \end{cases}$$

Note  $Y_1$  has  $E Y_1 = 0$  and  $Y_1 \in \{x_1, y_1\}$ . Similarly, for  $i \geq 2$ , given

$Y_{i-1} = y_i$  we have  $E Y_i = y_{i-1}$  and  $Y_i \in \{x_i, y_i\}$  and

given  $Y_{i-1} = x_j$  for some  $j \leq i-1$  we have  $Y_i = x_j$ . Finally note  $Y_{n-1} = X$ .

Hence the lemma will follow by showing  $(B(T_i), \dots, B(T_{n-1})) = (Y_i, \dots, Y_{n-1})$ .

This is immediate from the above discussion (since a RV taking two given values is determined by its expectation).

Lemma:  $T_{n-1}$  from the prev. lemma equals  $\tau$  of the Azéma-Yor thm.

Proof: Let  $j$  be such that  $B(T_{n-1}) = x_j$ . Note  $\psi(B(T_{n-1})) = y_{j-1}$ .

If  $j \leq n-1$  then  $T_{j-1} \leq T_j = \tau = T_{n-1}$  and  $B(T_{j-1}) = y_{j-1}$ .

If  $j = n$  then  $B(T_{n-1}) = x_n = y_{n-1}$ . Hence in both cases  $M(T_{n-1}) \geq y_{j-1}$ .

It follows that  $\tau \leq T_{n-1}$  since  $y_{j-1} = \psi(x_j) = \psi(B(T_{n-1}))$ .

Conversely if  $T_{i-1} \leq t \leq \tau$  for some  $i \leq j$  then  $B(t) \in (x_i, y_i)$

and we deduce  $M(t) < y_i \leq \psi(B(t))$ . Hence  $\tau \geq T_{n-1}$  establishing the lemma.

The rest of the proof is to use a limiting process to

Exercise

pass to general RVs. This is left as an exercise.

The Donsker invariance principle

Let  $(X_n)_{n \geq 1}$  be a sequence of IID RVs with  $E X_1 = 0$

and  $\text{Var}(X_1) = 1$ . Denote  $S_n = \sum_{i=1}^n X_i$ ,  $s(t)$  as its linear interpolation

$s(t) := s_{\lfloor t \rfloor} + (t - \lfloor t \rfloor)(s_{\lfloor t \rfloor + 1} - s_{\lfloor t \rfloor})$ . Note  $t \in \mathbb{R}$ . Now

define  $\{s_n^* | n \geq 1\}$  of random fcn. in  $\mathcal{C}([0, 1])$  by  $s_n^*(t) := \frac{s(t)}{\sqrt{n}}$ ,  $t \in [0, 1]$ .

Functional  
Central  
limit thm.  
or inv.  
principle

Thm.:  $s_n^*$  conv. in dist. to a std. BM  $B$  in the space  $\mathcal{C}([0, 1])$  (with sup norm).

Remark: This means that for any bdd. cont. function  $g$

on  $\mathcal{C}([0, 1])$  we have  $E g(s_n^*) \xrightarrow{n \rightarrow \infty} E g(B|_{[0, 1]})$ . Moreover, this holds

also for general bdd. meas.  $g$  with  $P(g \text{ is discont. at } B) = 0$ .

Porrmanteau  
thm.

It is also equivalent to having  $\forall$  closed  $K \subseteq \mathcal{C}([0, 1])$ ,  $\limsup_{n \rightarrow \infty} P(s_n^* \in K) \leq P(B|_{[0, 1]} \in K)$

Exercise

Embedding lemma 5.24. Proof of Donsker's inv. Principle.

Dist. of max. of walk up to n, LIL in HW.

Lecture 13 6.6.11

scribe

HW sol.?

Teaching survey

### Stochastic integration

Consider smp on  $Z, S_n$ . What is  $ds_n$ ? It is the random increments  $\pm 1$

Consider a gambling strategy  $R_n = \sum_{i=1}^n b_i X_i$ . Can write  $R_n = \sum_{i=1}^n b_i ds_i$  Unif. ind.

Here  $b_i \in \mathbb{Z}$  is a random predictable seq.,  $b_i \in \mathcal{O}(X_1, \dots, X_{i-1}, X_i \pm 1)$  Unif. ind.

Thus  $R_n$  is the money gained by time n and  $\mathbb{E}R_n = 0$  is a martingale.

We would similarly want to define  $\int_0^t H(s) dB(s)$ .

However,  $B$  is nowhere diff, hence of unbd. Variation, cannot define a Stieltjes integral.

We will see that for suitable integrands we will be able to define the integral as a random process.

Denote  $(\Omega, \mathcal{A}, \mathbb{P})$  our prob. space.

Suppose  $\{\mathcal{F}(t)\}_{t \geq 0}$  is a filtration

to which our BM is adapted, such

that the SMP holds, and which is

$(\forall A \in \mathcal{A}, \mathbb{P}(A) = 0 \text{ we have } A \in \mathcal{F}(t) \forall t)$ .

E.g.,  $\mathcal{F} = \text{completion of } \mathcal{F}^+$ .

Def:  $\{X(t, \omega) | t \geq 0, \omega \in \Omega\}$  is progressively measurable

if  $\forall t \geq 0$ , the mapping  $X: [0, t] \times \Omega \rightarrow \mathbb{R}$  is measurable

wrt.  $\mathcal{B}([0, t]) \otimes \mathcal{F}(t)$  borel sets

Remark: such  $X$  are in particular adapted to  $\mathcal{F}(t)$ .  $(X(t) \in \mathcal{F}(t) \forall t)$

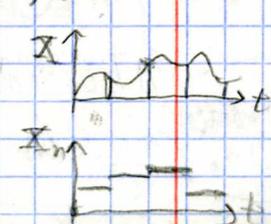
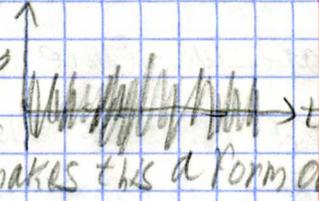
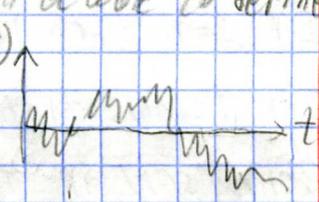
Lemma: Any  $X$  which is adapted and either left or right cont. is prog. meas. WLOG

Proof: Assume  $X$  is right-cont. Fix  $t > 0$ . For a positive int. n and  $0 \leq s \leq t$ , let  $X_n(s, \omega) = X(s, \omega)$  and  $X_n(s, \omega) = X(\frac{s+t}{2^n}, \omega)$   $kt \leq s < (k+1)t/2^n$

$(s, \omega) \mapsto X_n(s, \omega)$  is  $\mathcal{B}([0, t]) \times \mathcal{F}(t)$  meas. by adaptedness.

By right-cont.,  $\lim_{n \rightarrow \infty} X_n(s, \omega) = X(s, \omega) \forall s \in [0, t], \omega \in \Omega$ .

Hence  $(s, \omega) \mapsto X(s, \omega)$  is  $\mathcal{B}([0, t]) \times \mathcal{F}(t)$  meas.



We will define  $\int_0^t H(s) dB(s)$  for  $H$  prog, meas. s.t.  $\|H\|_2^2 := \mathbb{E} \int_0^\infty H(s)^2 ds < \infty$ .  
 Then  $\int_0^t H(s) dB(s) := \int_0^t H(s) 1_{[0,t]}(s) dB(s)$ . This is far from the most general def. possible and other int. are useful too! However, it has the advantages that it is an isometry in  $L_2$ :  $\mathbb{E} \left[ \left( \int_0^\infty H(s) dB(s) \right)^2 \right] = \|H\|_2^2$  and that  $\left\{ \int_0^t H(s) dB(s) \mid t \geq 0 \right\}$  is a cont. mart. wrt.  $\mathcal{F}(t)$ .

$\hat{I}_0$  integral  
 Perhaps mention also Stratonovich.

Finally, it is easier to define and already useful.

We begin by defining for  $H(t, \omega) := \sum_{i=1}^K A_i(\omega) 1_{(t_i, t_{i+1}]}(t)$   $0 \leq t_1 \leq \dots \leq t_{K+1}$  and  $A_i \in \mathcal{F}(t_i)$  (predictable) prog. meas.  
 $\int_0^\infty H(s) dB(s) := \sum_{i=1}^K A_i (B(t_{i+1}) - B(t_i))$

Now, we will show that a prog, meas.  $H$  with  $\|H\|_2^2 < \infty$  can be approx. in  $L_2$  by prog, meas. step proc.  $H_n$ . We then define  $\int_0^\infty H(s) dB(s) = L_2\text{-}\lim_{n \rightarrow \infty} \int_0^\infty H_n(s) dB(s)$  and show the limit exists and is ind. of the approx. seq.

Lemma: For a prog, meas.  $H$  with  $\|H\|_2^2 < \infty$  there exists a seq.  $\{H_n\}_{n \geq 0}$  of prog, meas. step proc. s.t.  $\|H - H_n\|_2 \xrightarrow{n \rightarrow \infty} 0$ .

Proof: We approx.  $H$  by 1) a bdd. prog, meas. proc.  
 2) a bdd. a.s. cont. prog, meas. proc.  
 3) a prog, meas. step process.

p. 192, 19, truncation only from above

1) For  $n > 0$ , let  $H_n(s, \omega) := H(s, \omega) 1_{[0, n]}(s)$ . Then  $H_n \xrightarrow[n \rightarrow \infty]{L_2} H_0$ .  
 For  $m > 0$ , we let  $H_{n,m}(s, \omega) := \begin{cases} H(s, \omega) & |H(s, \omega)| \leq m \\ 0 & \text{o/w} \end{cases}$ .  
 Then  $H_{n,m} \xrightarrow[m \rightarrow \infty]{L_2} H_n$ .

2) Assume  $H$  is unif. bdd. and on a compact int. Let  $h = \frac{1}{n}$  and, denoting  $H(s, \omega) := H(s, \omega)$  for  $s < 0$ , let  $H_n(s, \omega) := \frac{1}{n} \int_{s-h}^s H(t, \omega) dt$ .

$H_n$  is still prog, meas. (int. over past) and is a.s. cont.

Since  $\lim_{h \downarrow 0} H_n(s, \omega) = H(s, \omega)$  for a.e.  $s$ , a.s., we obtain

from the bdd. conv. thm. that  $H_n \xrightarrow[n \rightarrow \infty]{L_2} H_0$ .

3) Assume  $H$  is unif. bdd., a.s. cont. and on a compact int. Let  $H_n(s, \omega) := H(\frac{j}{n}, \omega)$  for  $\frac{j}{n} \leq s < \frac{j+1}{n}$ ,  $H_n$  is prog, meas. and  $H_n \xrightarrow[n \rightarrow \infty]{L_2} H$  again by bdd. conv.

Lemma: For a prog, meas. step process  $H$  with  $\|H\|_2^2 < \infty$  we have  $\mathbb{E} \left[ \left( \int_0^\infty H(s) dB(s) \right)^2 \right] = \|H\|_2^2$ .

Proof: Write  $H = \sum_{i=1}^K A_i 1_{(a_i, a_{i+1}]}$ .  $\mathbb{E} \left[ \left( \int_0^\infty H(s) dB(s) \right)^2 \right] = \mathbb{E} \sum_{i,j=1}^K A_i A_j (B(a_{i+1}) - B(a_i)) (B(a_{j+1}) - B(a_j)) =$   
 $= \sum_{i=1}^K \sum_{j=1}^K \mathbb{E} A_i A_j (B(a_{i+1}) - B(a_i)) (B(a_{j+1}) - B(a_j)) + \sum_{i=1}^K \mathbb{E} A_i^2 (B(a_{i+1}) - B(a_i))^2 = \sum_{i=1}^K \mathbb{E} A_i^2 (a_{i+1} - a_i) = \|H\|_2^2$   
 = 0 by the martingale prop. and pred. predictability

Cor: IF  $\{H_n\}_{n \geq 0}$  are prog. meas. step fns. with  $\|H_n - H_m\|_2 \xrightarrow{n, m \rightarrow \infty} 0$   
 then  $\mathbb{E} \left[ \left( \int_0^t (H_n(s) - H_m(s)) dB(s) \right)^2 \right] \xrightarrow{n, m \rightarrow \infty} 0$ .

Thm: IF  $H, \{H_n\}_{n \geq 0}$  are prog. meas.,  $H_n$  step proc.,  $H_n \xrightarrow{L_2} H$   
 then  $\lim_{n \rightarrow \infty} \int_0^t H_n(s) dB(s) =: \int_0^t H(s) dB(s)$  exists in  $L_2$ , is ind.  
 of the choice of  $\{H_n\}_{n \geq 0}$  and  $\mathbb{E} \left( \int_0^t H(s) dB(s) \right)^2 = \|H\|_2^2$ .

Proof: Immediate conseq. of Lemma, Cor. and the compo. of  $L_2$ .

Remark: IF  $H_n \xrightarrow{L_2} H$  surf. fast, that is,  $\sum_{n=1}^{\infty} \mathbb{E} \int_0^{\infty} (H_n(s) - H(s))^2 ds < \infty$   
 then  $\sum_{n=1}^{\infty} \mathbb{E} \left( \int_0^{\infty} (H_n(s) - H(s)) dB(s) \right)^2 < \infty$  and hence a.s.,  $\sum_{n=1}^{\infty} \left( \int_0^{\infty} H_n dB - \int_0^{\infty} H dB \right)^2 < \infty$   
 and hence  $\int_0^{\infty} H_n(s) dB(s) \xrightarrow{n \rightarrow \infty} \int_0^{\infty} H(s) dB(s)$ .

Def: For a prog. meas.  $H$  with  $\mathbb{E} \int_0^t H(s, \omega)^2 ds < \infty$  we let  $\int_0^t H(s) dB(s) := \int_0^t H(s) dB(s, \omega)$

Def: We say that a stoch. proc.  $X$  is a modification of the stoch. proc.  $Y$  if  $P(X(t) = Y(t)) = 1 \forall t$ .

Thm: IF  $H$  is prog. meas. and  $\mathbb{E} \int_0^t H(s, \omega)^2 ds < \infty \forall t \geq 0$  then there is an a.s. cont. modification of  $\left\{ \int_0^t H(s) dB(s) \right\}_{t \geq 0}$ . Moreover this process is a mart. and hence, in part,  $\mathbb{E} \int_0^t H(s) dB(s) = 0 \forall t \geq 0$ .

We will need Doob's <sup>LP</sup> maximal inequality: For a cont. mart.  $X$  and  $p > 1$  then for any  $t \geq 0$ ,  $\mathbb{E} \left( \sup_{0 \leq s \leq t} |X(s)|^p \right) \leq \left( \frac{p}{p-1} \right)^p \mathbb{E} |X(t)|^p$ .

Proof by the discrete version and approx.

Proof: Fix some  $t_0$  and let  $\{H_n\}_{n \geq 0}$  be step proc s.t.  $\|H_n - H\|_{L^2, t_0} \rightarrow 0$

Then  $\mathbb{E} \left( \int_0^{t_0} (H_n(s) - H(s)) dB(s) \right)^2 \xrightarrow{n \rightarrow \infty} 0$  (\*).

Since  $\forall s \leq t$ ,  $\int_0^s H_n(\omega) dB(\omega)$  is  $\mathcal{F}(s)$  meas. and  $\int_0^s H_n(\omega) dB(\omega)$  is ind. of  $\mathcal{F}(s)$ , then  $\left\{ \int_0^t H_n(\omega) dB(\omega) \right\}_{0 \leq t \leq t_0}$  is a martingale,  $\forall n$ .

Define  $X(t) := \mathbb{E} \left( \int_0^t H(s) dB(s) \mid \mathcal{F}(t) \right)$  for  $0 \leq t \leq t_0$ .  $X$  is also a mart. and  $X(t_0) = \int_0^{t_0} H(s) dB(s)$ . By the  $L_2$  max. ineq.:

$\mathbb{E} \left( \sup_{0 \leq t \leq t_0} \left( \int_0^t H_n(s) dB(s) - X(t) \right)^2 \right) \leq 4 \mathbb{E} \left[ \left( \int_0^{t_0} (H_n(s) - H(s)) dB(s) \right)^2 \right] \xrightarrow{\text{unif. by (*)}} 0$ .

By taking a subseq. we see that  $X$  is the a.s. limit of cont. process and hence is a.s. cont. By taking  $L_2$  limit  $\int_0^t H(s) dB(s)$  is  $\mathcal{F}(t)$  meas. and  $\int_0^t H(s) dB(s)$  is ind. of  $\mathcal{F}(t)$  with zero expect.

Hence  $X(t) = \int_0^t H(s) dB(s)$ .

## Ito's Formula

Fundamental thm. of Calc.:  $f(x(t)) - f(x(0)) = \int_0^t f'(x(s)) dx(s)$

When  $x$  is cont. and of bounded variation.

Thm. (Ito's Formula I): For a  $C_2$   $f: \mathbb{R} \rightarrow \mathbb{R}$  with  $\mathbb{E} \int_0^t f'(B(s))^2 ds < \infty$  for some  $t > 0$

We have a.s. for all  $0 \leq s \leq t$ :

$$f(B(s)) - f(B(0)) = \int_0^s f'(B(u)) dB(u) + \frac{1}{2} \int_0^s f''(B(u)) du.$$

Thm. (Ito's Formula II): Suppose  $\{s(s) | s \geq 0\}$  is an <sup>a.s.</sup> increasing cont. proc.

Let  $f: \mathbb{R}^2 \rightarrow \mathbb{R}$  be  $C_{2,1}$ . Assume  $\mathbb{E} \int_0^t [\partial_x f(B(s), s(s))]^2 ds < \infty$

for some  $t > 0$ . Then a.s. for all  $0 \leq s \leq t$ ,

$$f(B(s), s(s)) - f(B(0), s(0)) = \int_0^s \partial_x f(B(u), s(u)) dB(u) + \int_0^s \partial_y f(B(u), s(u)) ds(u) + \frac{1}{2} \int_0^s \partial_{xx} f(B(u), s(u)) du.$$

Multidimensional Ito's Formula:  $f: \mathbb{R}^{d+m} \rightarrow \mathbb{R}$   $C_{2,2,\dots,2,1}$ . Denote

$$\nabla_x f = (\partial_1 f, \dots, \partial_d f), \quad \nabla_y f = (\partial_{d+1} f, \dots, \partial_{d+m} f),$$

$$\int_0^t \nabla_x f(B(u), s(u)) \cdot dB(u) := \sum_{i=1}^d \int_0^t \partial_i f(B(u), s(u)) dB_i(u)$$

$$\int_0^t \nabla_y f(B(u), s(u)) \cdot ds(u) = \sum_{i=1}^m \int_0^t \partial_{d+i} f(B(u), s(u)) ds_i(u)$$

and  $\Delta_x f = \sum_{j=1}^d \partial_{jj} f$ . Then if  $B$  is a  $d$ -dim. BM,  $\{s(s)\}_{s \geq 0}$  is cont.

adapted with values in  $\mathbb{R}^m$  and increasing comp., and for some  $t > 0$

$\mathbb{E} \int_0^t |\nabla_x f(B(u), s(u))|^2 ds < \infty$  then a.s.  $\forall 0 \leq s \leq t$ :

$$f(B(s), s(s)) - f(B(0), s(0)) = \int_0^s \nabla_x f(B(u), s(u)) \cdot dB(u) + \int_0^s \nabla_y f(B(u), s(u)) \cdot ds(u) + \frac{1}{2} \int_0^s \Delta_x f(B(u), s(u)) du.$$

Remark: Since the formula holds a.s.  $\forall s \in [0, t]$ , it holds also

at stop. times  $\text{hd. by } t$ . In part, if  $f: U \rightarrow \mathbb{R}$  satisfies

the diff. cond. on an open set  $U$  and  $K \subset U$  is compact then

we may replace  $f$  by  $f^* = f \cdot g$ , where  $g$  is smooth with compact

supp. in  $U$ ,  $g \equiv 1$  on  $K$ . Let  $T$  be the first exit time from  $K$ .

It follows that the thm. applies to  $f^*$ , and hence to  $f$ , at times

$s \wedge T$  for all  $s < \infty$  simultaneously.